			Measurement Instrum	ent	Psychometric Characteristics	
Reference	Sample	Name of the Scale	Domains and Constructs	Length and Format of Instrument	Validity	Reliability
Kinney & Stephens (1989) <sup>2</sup> United States		Caregiver Hassles Scale (CHS)	Stress or hassles of daily living Five domains: (1) Assisting with ADLs; (2) Assisting with IADLs; (3) Cognitive status of patient; (4) Behavior of patient; (5) Social network of CG	wasn't, 2=Somewhat,	and the total score on the assigned "domain" or subscale (minus the item). Items with weak correlations were dropped resulting in a reduction from an initial pool of 110 item to 42 items. <u>Concurrent validity</u> was assessed by significant Pearson correlations between (a) the CHS-ADL subscale and the London Psychogeriatric Rating Scale (LPRS) measures of physical limitations (r=0.44, p<.001), and (b) the CHS-behavior hassles subscale and the LPRS-irresponsible behavior (r=0.331, p<.02). The CHS-cognitive status of patient subscale did not correlate significantly with the LPRS measure of "cognitive confusion."	Cronbach's $\alpha$ estimate, full scale=0.91 Cronbach's $\alpha$ by subscales:  ADL (Cronbach's $\alpha$ =0.79) Instrumental ADL (Cronbach's $\alpha$ =0.75) Cognitive (Cronbach's $\alpha$ =0.82) Behavior (Cronbach's $\alpha$ =0.89) Social network (Cronbach's $\alpha$ =0.74) Test-retest reliability (1-day interval, N=60) was estimated with Pearson's correlations. The reliability coefficient for the full scale=0.83 Test-retest reliability by subscales: ADL=0.86; IADL=0.71; Cognitive=0.80; Behavior=0.87; Social network=0.66
Lawton <i>et al.</i> (1989) <sup>3</sup> United States		Caregiver Appraisal Scale (CAS)	Appraisal of caregiving stress Three factors: (1) Subjective burden; (2) Caregiving impact; (3) Caregiving satisfaction	19 items, 5-point Likert scale (ranging from 1=Never True to 5=Nearly Always True <u>or</u> 1=Strongly Disagree to 5=Strongly Agree)	The <u>structural validity</u> of CAS was evaluated first with PCA using two independent samples and secondly through a CFA. (The first independent sample reported here (N=632) consisted of AD caregivers. The second cross-validation sample comprised a mixed sample of CGs.) The results of the PCAs with the two independent samples were used to refine the original 47-item scale with 5 components/factors resulting in a reduced 19-item scale with 3 factors.  The CFA was conducted with the same two independent samples confirming an underlying 3-factor structure. The first sample (N=632) yielded acceptable fit indexes (e.g., GFI=0.94, NFI=0.90). Results in the cross-validation sample were lower (e.g., GFI=0.86, NFI=0.78).  Concurrent validity was established through correlations of the three CAS subscales with the following measures: Burden rating; Quality of relationship; Emotional burden; Relationship to impaired person. Subjective burden was highly related to Burden rating scores (r=0.65) and less strongly but significantly to all of the other scales (r's=0.28-0.33). Caregiving satisfaction was less strongly related to the Burden rating (r=0.24) but strongly related to the quality of the relationship to the impaired person (r=0.50). Caregiving impact was highly correlated with Burden rating (r=0.57).	Subjective burden (α=0.85)
Ellis et al. (1989) <sup>5</sup> United States		(CRS)	Reactions to caregiving Seven factors/dimensions: (1) Financial impact; (2) Impact on schedule; (3) Restrictions in social activities; (4) Impact on health; (5) Caregiving role responsibility; (6) Negative reactions; (7) Family abandonment of CG Perceptions of availability of social resources One factor: Availability of resources	34 items, 5-point Likert scale (ranging from 1=Strongly disagree to 5=Strongly agree)  6 items, 5-point Likert scale (ranging from 0=No assistance to 4=Most frequent amount of assistance)	Authors did not explicitly talk about <u>content validity</u> of the items in the scales developed but report conducting a review of the literature to define the concepts included in each of the scales and key relationships between concepts that needed to be considered in the development of the scales. In particular, the development of the "Caregiver Reactions" scale included in-depth interviews with CGs of persons with various types of physical and cognitive impairments. A pool of 101 items were identified from both the literature review and the analysis of the interviews.  The <u>structural validity</u> of the CRS was established through a CFA to test a theorized 7-factor structure. One of the hypothesized subscales ("restrictions in social activities") was dropped from the final solution as well as items from the original pool. The final scale consisted of 34 items and 6 factors. No GFI statistics are reported for the CFA model.  The <u>structural validity</u> of the SRS was established through CFA to test a theorized one-factor structure. The single factor produced factor loadings with acceptable ranges (0.42-0.62). No GFI statistics are reported for the CFA model.	Cronbach's $\alpha$ by subscales: Financial impact of caregiving ( $\alpha$ =0.77) Impact on schedule ( $\alpha$ =0.84) Impact on health ( $\alpha$ =0.81) Caregiving role responsibility ( $\alpha$ =0.88) Negative reactions to caregiving ( $\alpha$ =0.83) Family abandonment of CG ( $\alpha$ =0.87)  Cronbach's $\alpha$ , full scale=0.69
Kosberg et al.(1990) <sup>6</sup> United States		The Cost of Care Index (CCI)	Consequences (or costs) of caretaking Five factors/components: (1) Personal and Social Restriction; (2) Physical and Emotional Health; (3) Value of Providing Care; (4) Patient as Provocateur; (5) Economic Costs	20 items, 4-point Likert scale (ranging from 1=Strongly agree to 4=Strongly	Content validity was established by first reviewing the literature related to the "costs" of providing care to dependent elderly persons and defining the dimensions needed to develop a bank of items. Twenty-seven items were initially derived from the input of professionals working with family CGs of frail and impaired elderly relatives. These 27 items were pilot tested to determine their ability to distinguish groups of carers caring for elders with different physical and mental impairment levels.  The initial inspection of the structural validity of the CCI was not conducted in a sample of dementia CGs. Instead, authors recruited a sample of 137 CGs of clients seeking nursing home placement under the Florida Medicaid Program. A PCA with Varimax rotation yielded a final 20-item scale with 5-components/factors. The factorial structure of the CCI scale was not examined in the present study of dementia CGs.	<u>Cronbach's α, full scale</u> =0.79

						Concurrent validity was demonstrated by statistically significant Pearson's correlations between CCI	
						scores and, for example, measures of caregiving functioning assessed by a) the Short Psychiatric	
						Evaluation Schedule (SPES; r=0.27, p<0.01), b) self-reported mental health (r=0.36, p<0.001), and c)	
						physical health (r=0.22, p <0.05). Significant correlations were also obtained between CCI scores and	
						measures of "consequences of caregiving" assessed by a) the "ADL trouble due to patient" (0.24, p <	
						0.01) and b) "Tolerance for patient behavior" (-0.33, $p < 0.001$ ).	
	dman	ADRD	Perceived Social	Perceived social support	9 items,	Content validity was demonstrated by a team of researchers writing an initial 21-item pool based on a	Cronbach's α, full scale=0.84
(199	1) <sup>7</sup>		Support for	One factor: Availability or	5-point Likert scale	review of literature and empirically-determined reasons for joining self-help groups.	
			Caregiving	adequacy of social support	(ranging from 1=Not at all	Structural validity was established through PCA with Varimax rotation using an initial 21-item scale.	
Unit	ed States		(PSSC)	(or help)	to 5=Extremely)	Inspection of item loading reduced the scale to 12 items. PCA analyses were repeated yielding two	
						separate scales (each with one factor/component): The Perceived Support for Caregiving (PSSC) and the	
						Social Conflict (SC) scale. The PSSC explained 42.8% of the variance.	
						Concurrent validity was established by significant positive Pearson correlations between PSSC total	
						scores and Natural Network Indices (r=0.26 to 0.39; p<0.02 to 0.001). These results were collected from	
						a subsample of respondents (N=70 to 79).	
			Social Conflict	Social conflict	3 items,	Structural validity. A PCA with Varimax rotation yielded the SC factor/component explaining 18% of the	Cronbach's α, full scale =0.72
			(SC)	One factor: Lack or	5-point Likert scale	variance.	<u> </u>
			` '			Concurrent validity was established by a significant positive Pearson correlation between the total	
				(or help)	to 5=Extremely)	scores on the SC and the ZBI item "Do you feel that your relative currently affects your relationship with	
				(5)		other family members or friends in a negative way" (r=0.34, $p$ =0.001).	
Thei	ut et al.	ADRD	Anticipatory	Anticipatory grief	27 items,		Cronbach's α, full scale =0.84
(199				(bereavement of wives	5-point Likert scale	the dimensions of anticipatory grief and developing additional items based on clinical experience with	erombaerr 5 a, rain Searc
(133	-,			l'		wives of patients with dementia.	
Unit	ed States				, , , ,	Structural validity. No formal examination of the underlying structure of the scale is presented.	
Office	cu states				agree, 4=Agree,	Concurrent validity was established by positive and significant (p<0.001) correlations between total	
						scores in the AGS scale and the depression, anxiety, and hostility dimensions of the Hopkins Symptom	
				Anxiety; (4) Irritability; (5)	D-Strongly agree	Checklist (SCL-90-R). Note: Coefficients were not reported.	
				Sadness; (6) Feelings of loss;		eneckist (See 30 N). Note: Coemicins were not reported.	
				(7) Decreased function			
\/i+al	liano et	ADRD		Burden or distress	25 items,	Content validity was established by a review of extant literature on problems in AD caregiving and by	Cronbach's α by subscales:
	1991) <sup>9</sup>	ADIO		Two domains:	1		Objective burden (α=0.85)
ui. (1	1331)						Subjective burden (α=0.89)
Lloit	ed States		, ,	number of negative	1=Occurrence)		Test-retest reliability (15-18 months
Office	eu states			_	· · · · · · · · · · · · · · · · · · ·		apart) was estimated with Pearson's
					point scale (ranging from		correlations between scale
					0=No occurrence to		administrations:
				'		(0.54, 0.41, respectively), anxiety (0.43, 0.26, respectively), and suppressed anger (subjective=0.42). The	
							Subjective burden (r=0.04, p<0.001)
						respectively). Depression, anxiety, suppressed anger, and morale were measured by the Beck	Subjective burden (1-0.70, p<0.001)
					l <sup>*</sup>	Depression Inventory-Short (BDI-S), the Symptom Checklist90 Anxiety Scale (SCL-90), the Suppressed	
					_	Anger Subscale from the Anger Expression Scale, and the Satisfaction with Life Scale, respectively.	
					subjective burden	Anger Subscale from the Anger Expression Scale, and the Satisfaction with the Scale, respectively.	
Givo	n et al.	Mixed	Caregiver		24 items;	Content validity was established by requiring that all the members of a 5-group team agree that each	Cronbach's α by subscales:
(199		iviiAeu	_		-	item from a 101-item pool belonged to a particular construct. After a pilot test with 99 CGs, items were	
(133						dropped due to lack of clarity, variability, or association with any other items reducing the item pool to	
Unit	ed States				to 5=Strongly agree)		Impact on schedule (α=0.82) Impact on finances (α=0.81)
Office	eu states		, ,	Impact on CG's daily	to 3-3trongry agree)		Sense of self-esteem ( $\alpha$ =0.90)
				schedule; (3) Impact on CG's			Friends/family support ( $\alpha$ =0.85)
				finances; (4) Relationship to		solution. A scree plot confirmed the 5-factor structure. Further elimination of items due to low loadings	Therias, fairing support (a=0.83)
				CG's sense of self-worth; (5)		resulted in a 24-item scale. A final re-run of the PCA model with the 24-item scale yielded a 5-factor	
				Friends/family support		solution accounting for 65.1% of the variance in items.	
				i rienus/ranniy support		Measurement (factorial) invariance tests. Using a second independent sample (also N=377; 26.8%	
						dementia CGs) a multiple-group CFA was conducted to test the factorial invariance of the CRA	
						instrument across diseases (Alzheimer's vs. cancer), spouse vs non-spouses, and across time. Results	
						showed that the subscale structure of the CRA remains stable across diverse groups of CGs and across	
						time. The latter suggests suitability of the CRA to measure change in CG reactions.	
						, ,	
						Concurrent validity was established by calculating Pearson correlations between CRA subscales and a)	

Semple (1992) <sup>11</sup> United States	Scales (FCS)	Family conflict Three factors: (1) Definitions & strategies conflict; (2) Treatment of patient conflict; (3) Treatment of CG conflict	12 items, 4-point Likert scale (ranging from 1=No disagreement to 4=Quite a bit of disagreement)	the CES-D and b) the ADL Dependencies Scale using the combined independent samples (N=754). Finance (r=0.25, r=0.34), family support (r=0.20, r=0.39), and health (r=0.29, r=0.57) were significantly and positively correlated with ADL dependencies and depression respectively. CG esteem was significantly and negatively associated with depression (r=-0.23).  Content validity was established by in-depth interviews with 20 CGs that resulted in the identification of three family conflict domains and the creation of 4 items per domain.  Structural validity was established through CFA using the 12-item scale. After comparing competing model that conformed underlying theories, a 3-factor model yielded the best fit as measured by a GFI=0.98 and a chi-square/df ratio=2.9 (less than 3 is desirable).  As evidence of concurrent validity the authors used Pearson correlations to show relationships between the three FCS factors/subscales and the Hopkins Symptom Checklist (HSC) measures of anger and depression. All correlations between the HSC-Anger measure and the FCS subscales were significant (p<0.001): Definitions & strategies conflict (r=0.23), Treatment of patient conflict (r=0.25), and Treatment of CG conflict (r=0.34). All correlations between the HSC-depression measure and the FCS subscales were also significant (p<0.001): Definitions & strategies conflict (r=0.23), Treatment of patient conflict (r=0.24).	Definitions & strategies conflict $(\alpha$ =0.80) Treatment of patient conflict (=0.86) Treatment of CG conflict (reported as
Teri <i>et al.</i> (1992) <sup>12</sup>	•	CG reactions to patient behavior problems		conflict (r=0.28), and Treatment of CG conflict (r=0.28).  Content validity was shown by raters sorting items into hypothesized content areas, rating items, and agreeing on items. This method reduced the original pool of 64 items to 47 items.  A BCA approach with Various vectoring was used to study the underlying dimensions of the scale using	Behavior <i>Frequency</i> Scoring: <u>Cronbach's α, full scale</u> =0.84 <u>Cronbach's α by subscales:</u>
United States	(RMPBC)	problems; (2) Depression problems; (3) Disruptive behaviors (The scale uses two scoring methods per item: frequency of patient behavior problems and CG distress or reaction to the patient behavior problems.)	in the past week, 2=1 to 2 times in the past week, 3=3 to 6 times in the past week, 4=Daily or more often) 2) Reaction of "upset" by CG: 5-point Likert scale (0=Not at all, 1=A little, 2=Moderately, 3=Very much, 4=Extremely)	Related Problems (r=0.29), Depression (r=0.31), and Disruption (r=0.26) subscales. Burden measured by the CSS correlated with the RMPBC Memory-Related Problems (r=0.32), Depression (r=0.42), and Disruption (r=0.41) subscales. <u>Discriminant validity</u> was established for RMPBC Behavior Frequency by non-significant correlations between the RMPBC Depression subscale and the Mini-Mental State Exam (r=-0.04, p>0.05) as well as non-significant correlations between the RMPBC Memory-Related problems subscale and the HAM-D Scale (r=0.001, p>0.05).	Depression ( $\alpha$ =0.80)  Memory-Related problems ( $\alpha$ =0.79)  Disruption ( $\alpha$ =0.67)  Caregiver <i>Reaction</i> Scoring: <u>Cronbach's <math>\alpha</math>, full scale</u> =0.90 <u>Cronbach's <math>\alpha</math> by subscales:</u> Depression ( $\alpha$ =0.89)  Memory-Related problems ( $\alpha$ =0.88)  Disruption ( $\alpha$ =0.84)
Macera <i>et al.</i> (1993) <sup>13</sup> United States	Caregiver Burden Scale (CBS)	Perceived burden Three domains: (1) Activity for which patient required help; (2) Activity for which CG provided help; (3) Stress by providing help	15 items, 2-point scale (0=No, 1=Yes)	Content validity and the creation of items was not addressed in the article. Authors reviewed the literature on perceived burden and state the importance of measuring burden associated with specific caregiving tasks. Results of the authors-developed CBS scale are presented as a pilot study.   Structural validity. No examination of the underlying structure of the scale is presented.  Concurrent validity for the CBS scale was established by a significant positive Pearson correlation with the CES-D (r=0.38, p<0.001).	<u>Cronbach's α, full scale</u> =0.87
Hinrichsen & Niederehe (1994) <sup>16</sup>		Encouragement; (3) Active		pilot interviews. <u>Structural validity</u> was established through EFA using PAF extraction and Varimax rotation that yielded a	Cronbach's $\alpha$ by subscales: Criticism ( $\alpha$ =0.85) Encouragement ( $\alpha$ =0.80) Active management ( $\alpha$ =0.77)
United States		management	5=Most of the time)	3-factor solution. The original 34-item pool was reduced to 28 items based on factor loadings.	
Carruth (1996) <sup>18</sup> United States	Reciprocity Scale (CRS)	(1) Warmth and regard; (2) Intrinsic rewards for giving;	26 items, 5-point Likert scale (ranging from 1=Strongly disagree to 5=Strongly agree)	Content validity was established by an initial 50-item pool developed from a literature review and interviews with family CGs. Two panels of experts rated items relevance and CVIs were computed. Items with low CVIs were eliminated reducing the pool to 32 items. A pilot test with 30 CGs provided data for further item reduction by "item-to-item", "item-to-subscale", and "item-to-total" correlations further reducing the scale to 30 items.  Structural validity. Before attempting to establish validity, an inter-item analysis dropped four poorly-correlated items reduce the 30-item scale to 26 items. The sample (N=303) was randomly split into two subsamples to perform EFA (N=130) and CFA (N=173). An EFA performed by factor analysis with Varimax rotation yielded a 22-item, 4-factor solution that accounted for 62.9% of the variance. The CFA with the cross validation sample established the acceptability of the 4-factor model with adequate fit	Intrinsic rewards for giving ( $\alpha$ =0.82) Love and affection ( $\alpha$ =0.86) Balance within family caregiving ( $\alpha$ =0.78) Test-retest reliability was estimated using Pearson's correlations with a

					indexes (e.g., GFI=0.88; AGFI=0.85; RMR=0.05; TLI=0.95).  AVE was used to assessed the <u>convergent validity</u> of the 4 factors extracted by CFA. AVE values ranged from 0.47 to 0.64. (Three of the four AVEs were slightly below the recommended threshold of 0.50.)	Test-retest reliability by subscales: Warmth and regard (r=0.70); Intrinsic rewards for giving ( r=0.69); Love and affection (r=0.88); Balance within family caregiving (r=0.58)
Davis et al. (1997) <sup>22</sup> United States	ADRD	Activity Survey (CAS)	activities (One "total score" measure) Scores were the hours and minutes engaged in the	(1) communicating; (2) using transportation (3) dressing; (4) eating (5) looking after one's	Content validity. Special efforts were made to find terms that could be used with a variety of populations in different cultures. Several versions of the scale were pilot-tested. in different cultural settings. Specialists reviewed the scale drafts to develop cultural and linguistic equivalents in several languages. Specific definitions were provided, along with examples of what was meant for the various categories of assistance. As a result, an initial pool of 13-items was developed and tested. The results of the first analysis led to the reduction of the original 13-item scale to a 6-item CAS scale.  The concurrent validity of CAS was established by significant (p-values < 0.05) Pearson correlations with the Alzheimer's Disease Assessment Scale Cognitive Subscale (ADAS-Cog) (r = 0.61), MMSE (r = -0.57) and Physical Self Maintenance Scale (PSMS) (r=0.43).	Test-retest reliability was established by retesting N=42 CGs within a 2-week interval (i.e., week 1 and week 3), and calculating the ICC. The ICC=0.85, p<.001.
Picot <i>et al.</i> (1997) <sup>23</sup> United States	ADRD	Rewards Scale (Picot-CRS)	Two domains/subscales: (1) External rewards:	2=Somewhat, 3=Quite a lot, 4=A great deal)	Content validity was established by interviews with eight family CGs to identify themes about positive feelings and changes (i.e., rewards) resulting from caregiving. Twenty-seven items were generated from caregiving literature and considering caregiving's external and internal rewards. A pilot test with 20 CGs led to a reduction from 27 to 24 items. The underlying factorial structure of the scale was not examined. Concurrent validity was demonstrated by a significant positive Pearson correlation between PCRS scores and "caregiving demands" (r=0.22, p<0.05) measured by Texas Research Institute of Mental Sciences Behavioral Problem Checklist (TRIMS BPC) as well as by a significant positive association between PCRS scores and palliative coping (r=0.26, p<0.05) measured by the Jalowiec Coping Scale. A hypothesized negative association between rewards and costs as measured by the Costs of Care Index (CCI) was not found (r=0.07, p>0.05).	
Kaufer <i>et al.</i> (1998) <sup>25</sup> United States		The Neuropsychiatric Inventory Caregiver Distress (NPI-D) Scale	Subjective CG distress One domain: Psychological distress Note: The NPI-D was developed as an adjunct scale of Neuropsychiatric Symptoms (NPI) scale. The NPI-D assesses the impact of neuropsychiatric symptoms in Alzheimer's disease (AD)	0=Not at all distressing to 5=Very distressing) Note: The 10 items represent symptoms obtained from one of the three subscales from the original NPI scale. Items	Content validity. A preliminary version of the NPI-D included items from the three subscales contained in the original NPI scale (physical, social, and psychological distress). An initial field testing of the NPI-D scale revealed that AD CGs viewed the impact of neuropsychiatric disturbances primarily in terms of psychological or emotional distress. As a result, a revised version of the NPI-D scale excluded items from the physical and social subscales.  No examination of the underlying factorial structure of the NPI-D scale is presented.  Concurrent validity was established by correlating scores in the NPI-D scale in a subsample (N=69) with scores in an abridged version of the Relatives' Stress Scale (RSS) using Pearson's correlation (r=0.60, p<0.001). (The abridged RSS included 2 of the 3 subscales: personal distress and negative feelings.) The correlation between total NPI and NPI-D scores was 0.83 (p < 0.001).	correlation (r=0.92, p<.001). Interrater reliability was also calculated with the ICC between two raters of the NPI-D in <u>16 CGs</u>
Zeiss <i>et al.</i> (1999) <sup>26</sup> United States		Care Self-Efficacy	that reduce stress and enhance well being	performing item activity (ranging from 0%=No confidence to 100%=Completely confident).	Both measures were field tested in the same sample of 217 CGs. No examination of the underlying structure of the scales is presented. <u>Concurrent validity</u> for the Self-Care Self-Efficacy scale was established by a significant positive Pearson correlation between Self-Care and the "network size" subscale of the Arizona Social Support Interview (r=0.30, p<.001).	Test-retest reliability with a subsample (N=39) retested after 11 weeks was a high Pearson coefficient (r=0.675, p<0.001).
		Self-Efficacy	utilizing problem-solving skills shown to be positively related to psychological adjustment	Rating of confidence in		Cronbach's α, full scale =0.83.  Test-retest reliability with a subsample (18%) retested after 11 weeks was a high Pearson coefficient (r=0 .683, p<0.001).
Farran <i>et al.</i> (1999) <sup>27</sup> United States	ADRD	Through Caregiving Scale (FMTCS)	Positive aspects of caregiving Three factors: (1) Loss/Powerlessness (LP);	43 items, 5-point Likert scale (ranging from 1=Strongly disagree to 5=Strongly		$\frac{\text{Cronbach's }\alpha, \text{ full scale}}{\text{Cronbach's }\alpha \text{ by subscales:}} : \\ \text{Loss/Powerlessness (LP) } (\alpha = 0.89) \\ \text{Provisional Meaning (PM) } (\alpha = 0.88) \\ \text{Ultimate Meaning (UM) } (\alpha = 0.91)$

-						
	]		(PM); (3) Ultimate meaning		to-total correlations. This pilot produced "acceptable" reliability estimates both by three	
			(UM)		factors/subscales (0.88 to 0.95) and total scale (0.91). The pilot test-retest reliability (one-month	
					interval), estimated with Spearman correlation, ranged from 0.85-0.89 for the three subscales and 0.80	
					for the full FMTCS.	
					Given that the three original subscales had a "strong" theoretical base, the authors used CFA to	
					establish the factorial validity of the FMTCS using an independent sample of N=215 caretakers (only	
					N=208 had available data on the FMTCS). The CFA model confirmed the 3-factor structure identified in	
					the previous pilot study and provided an adequate overall fit (e.g., GFI=0.998 and a coefficient of	
					determination=0.763).	
					Concurrent validity was established by Pearson's correlations between FMTCS scores and existing	
					measures hypothesized to be related. Scores on the LP subscale were significantly (all p-values < 0.01)	
					correlated with scores on a) Patient Problem behaviors (r=0.44), b) Marital tension (r=0.38), c) Global	
					role strain (r=0.70), and d) depression (r=0.61), as measured by the CES-D. Scores on the PM subscale	
					were significantly correlated with a) Marital satisfaction (r=0.24), b) Caregiver Satisfaction (r=0.64), and	
					c) Personal gain(r=0.57). Scores on the UM subscale were significantly correlated with a) Religious	
					participation (r=0.53), b) Personal religion beliefs (r=0.61), and c) Religious support satisfaction (r=0.24).	
					Total FMTSC scores revealed similar relationships. FMTSC total scores were positively associated with	
					measures of a) Marital satisfaction (r=0.46), b) Caregiver satisfaction (r=0.58), c) Personal gain (r=0.39),	
					c) Religious participation (r=0.37), d) Religious beliefs (r=0.54), and e) Religious support satisfaction	
					(r=0.21). Total FMTSC scores, however, were <i>negatively associated</i> with Patient Problem behaviors (r=-	
1					0.35), Marital tension (r=-0.49), Role strain (-0.64), and Depression (-0.60).	
Gitlin <i>et al.</i>	ADRD	Task	CG strategies to simplify	19 items,	Content validity was demonstrated by gathering an initial 20-item pool from observational research and	Cronhach's α full scale in Sample 1
(2002)32	,			5-point Likert scale		(N=202) was 0.81.
(2002)			patients	· •	change the external environment by simplifying task requirements and interactions for the person with	(11 202) 1140 01021
United States			ľ'	5=Always)	, , , , , , , , , , , , , , , , , , , ,	Note: The Cronbach's α estimate for
otou otutes		, ,	cope with deficits in		,	the full scale in Sample 2 was slightly
			functioning, orientation, and		difficulties, a final 19-item TMSI included items that reflected constructive strategies that would benefit	
			awareness of patients			thresholds.
			awareness of patients		The <u>structural validity</u> of the 19-item TMSI was examined in an independent sample of 202 CGs (Sample	em estroras.
					1) using an EFA with principal-axis factoring extraction method. EFA identified one factor accounting for	
					60.2% of the variance in items. Factor loadings ranged from 0.35 to 0.87. Using an independent sample	
					of 255 CGs with similar characteristics as the sample used in the EFA, the <u>concurrent</u> validity was	
					demonstrated by computing Pearson's correlations between TMSI scores and a) functional dependency	
					of Only ADRD patients as measured by "ADL dependence" (0.237, p<0.001), b) CG self-efficacy, as	
					measured by "ADL self-efficacy" (0.173, 0<.05), and c) use of positive coping strategies, as measured by	
					a subscale of the Dementia Management Scale, DMS (0.507, p<0.001).	
					<u>Discriminant validity.</u> As expected, TMSI scores were not associated with a) level of CG upset with	
					disruptive behaviors, as measured by the Disruptive Behaviors subscale of the RMBPC (-0.002, p>.05)	
					and b) CG use of criticism-based strategies, as measured by a subscale of the DMS (-0.055, p>.05).	
Fortinsky <i>et</i>	ADRD	Family caregiver	Perceived Self-Efficacy	9 items,	Content validity was shown by reviewing the literature on perceived self-efficacy, applying its principles	Crophach's a by subscalos:
al. (2002) <sup>33</sup>	AUNU	, .	Two factors:	-		
ui. (2002)		,				Symptom management (α=0.77) Community support service use
United States			(2) Use of community	certain to 10=Very		community support service use (α=0.78)
Officed States			support services	certain to 10-very	The structural validity was demonstrated by an EFA with PAF as factor extraction method and Varimax	(4-0.78)
			support services	·	rotation to simplify factor interpretation. EFA identified two factors accounting for 54% of the variance.	
					A scree plot confirmed the two-factor solution. One item was eliminated due to low factor loading	
					resulting in a final 9-item scale.	
1					Concurrent validity was established by calculating Pearson's correlations between a "Global CG	
1					competence measure" and the two perceived self-efficacy factors: a) symptom management ( $r=0.49$ , p < 0.01) and b) use of community support services ( $r=0.27$ , p < 0.01).	
Marwit <i>et al.</i>	V D B D	Marwit–Meuser	CG griof	50 items,	, , , , , , , , , , , , , , , , , , , ,	Cronbach's α, full scale =0.96.
(2002) <sup>34</sup>	AUKU			50 Items, 5-point Likert scale	content validity was demonstrated by conducting 16 focus groups with N=90 dementia CGS. Focus groups resulted in the generation of a pool of 184 grief statements/items. A preliminary analysis of the	
(2002)5-						
United States			(1) Personal Sacrifice		·	Personal Sacrifice Burden (α=0.93)
United States		·	Burden; (2) Heartfelt	· · · · · · · · · · · · · · · · · · ·		Heartfelt Sadness and Longing
1				agree)	, , , , , , , , , , , , , , , , , , , ,	(α=0.90)
			Worry and Felt Isolation			Worry and Felt Isolation (α=0.91)
	L	<u> </u>	<u> </u>	<u> </u>	using the pool of 69 items, an EFA, with PAF extraction method and oblique rotation yielded a three-	Guttman's split-half estimate, full

				i T		factor solution explaining 34% of the item variance. Items with double loadings were dropped resulting	scale=0.91
			'	1	1	in a final 50-item MM-CGI scale.	Guttman's split-half by subscales:
				1	1	Concurrent validity was established by significant positive Spearman's rank correlations between scores	
			'	1	Í		split-half=0.91)
			'	1	Í		Heartfelt Sadness and Longing
			'	1			(Guttman's split-half=0.86)
				1		0.01). Statistically significant ( $p < 0.01$ ) negative correlations between scores of the MM-CGI and a) the	· · · · · · · · · · · · · · · · · · ·
				1			split-half=0.91)
				1	İ		Spiit-fiaii=0.51)
C: -ff	-1 -1		Carla fan	C -1 1 15 office ou		Questionnaire-Family subscale (rho=-0.36) also supported the convergent validity of this MM-CGI scale.	C. Laskia di Laskia di Laskia
	n et al.					To expand the <u>content validity</u> of the original self-efficacy measure developed by Zeiss et al. (1999), the	
(2002) <sup>3</sup>	33				Confidence in doing	authors conducted a thorough literature review and added 37 items mostly representing a new domain	
l							Responding to disruptive behavior
United	States					• • • • • • • • • • • • • • • • • • • •	(α=0.84); Controlling upsetting
				,		· · · · · · · · · · · · · · · · · · ·	thoughts (α=0.86)
				upsetting thoughts		examining the item distributions using responses from the first independent sample (N=169), the initial	
				1		- F	with a subset participants (N=100)
				1		·	after a 2-week interval using Pearson's
				1		, 9	correlation coefficients.
				1		1 ' ' ' ' ' ' ' ' ' ' ' ' ' ' ' ' ' ' '	Test-retest reliability by subscales:
				1	İ	15-item solution (e.g., CFI=0.93 and the $\chi$ 2/df = 1.59). (Of note, a value less than 3 is a commonly used	Obtaining respite (r=0.76); Responding
			'	1	Í	· · · · · · · · · · · · · · · · · · ·	to disruptive behavior (r=0.70);
				1	İ	The <u>concurrent validity</u> was established by significant Pearson's correlations between scores on	Controlling upsetting thoughts (r=0.76)
			'	1		different R-SCSE subscales (factors) and (a) depression, as measured by Short Form Beck Depression	
			'	1			Note: Reliability estimates by
			'	1	Í	Controlling upsetting thoughts: r=-0.38; p<0.001), (b) anger, measured by the Spielberger's Trait Anger	subscales were obtained in both
			'	1			independent samples. The pattern of
			'	1	Í	Anxiety (Controlling upsetting thoughts: r=-0.62, p<0.001), and (d) perceived social support network,	estimates was the same in the second
			<u> </u>				sample.
Mahon	ney <i>et l</i>	ADRD	Caregiver	Caregiver vigilance or	4 items,	Content validity was established through a year-long qualitative study collecting data from discussions	Cronbach's α, full scale =0.66.
al. (200	03) <sup>37</sup>		Vigilance Scale	oversight of patient	Items 1 and 2 are scored	with 70 family CGs on vigilance and oversight of care recipients. The study led to the key finding that a	
			(CVS)	activities	with two scales:	vigilant CG is actively involved and perceives herself as responsible for the care recipient even when not	
United	States			One factor: Oversight of	2-point, binary scale	"actively" providing care. As a consequence, four vigilant questions/items were developed that	
			'	patient activities (	(0=No, 1=Yes) and time	reflected "being there" and "doing things" for the care recipient. The items were pilot tested with 15	
			'	į l	F 5	family CGs resulting in the refinement and re-wording of questions.	
				1		A PCA was conducted to study the <u>structural validity</u> of the 4-item scale. The analysis yielded a single	
			'			component accounting for 50% of the variance.	
			'			Concurrent validity was supported by a significant negative Pearson's correlation between CVS scores	
						and scores on the MMSE (r=-0.34, p<0.001). The greater the cognitive impairment (lower MMSE score),	
			'			the greater the score in the CVS scale. The correlation between CVS scores and total scores on the	
			'	l F			
			'			(r=0.15, p<0.001).	
Steven	s et al.	ADRD	The Leisure Time		6 items,		Cronbach's α, full scale =0.80.
(2004)4			Satisfaction (LTS)		3-point Likert scale	measure of leisure that assessed the concept of satisfaction with leisure: the 51-item Leisure	Crombust 5 d, ran 35d.5
(200.,			' '	One factor: Satisfaction with	· ·	Satisfaction Scale (LSS). <sup>41</sup> However, this measure had not been evaluated with CGs of older adults and	
United	States			I I		had an estimated administration time judged inappropriate as a brief measure to assess changes in	
01	Juice			caregiving on leisure		leisure after caregiving interventions. Taken into account the review of literature and limitations of the	
				activities).		existing LSS tool, the authors developed a short 6-item scale to assess the distinct psychological	
				detivicies).		dimension of satisfaction with the amount of time spent in leisure activities relevant to family CGs of	
			'	1	Í	those with Alzheimer's disease or a related dementia.	
				1	İ	To establish the structural validity of the 6-item scale, the baseline sample (N=1225) with non-missing	
				1		item data was randomly split into two subsamples to perform a PCA (N=900-roughly 75% of the sample)	.
				1		and CFA (N=291-roughly 25% of the sample). A PCA, oblique rotation, and weighted least squares	
				1		estimation yielded a one-factor solution explaining 57.8% of the variance. The CFA indicated a good fit	
				1		for the one-factor solution with a RMSEA statistic of 0.069.	
				1			
			'	1		Concurrent validity was shown by "small to moderate" and significant (p<0.001) Spearman's rank	
			'	1	Í	correlations between scores on the LTS and a) a 3-item measure of <i>CG satisfaction with social support</i> (rho =0.32), b) <i>social network</i> , measured by the Lubben Social Network (rho =0.25), and c) wellbeing,	
			1 '	<u> </u>	1	(mo =0.32), b) social network, measured by the Lubben Social Network (mo =0.25), and c) wellbeing,	

		l			measured by the CES-D-wellbeing subscale (rho =0.28). Expected negative correlations with LTS scores	
					included time spent on ADL activities (rho =-0.21) and depression measured by the CES-D (rho =-0.37).	
Gaugler <i>et al.</i>				34 items,	Content validity was demonstrated by researchers identifying seven domains of unmet CG need from	Cronbach's α by subscales:
(2004)42		Need (PUN)	different "stages" of the	2-point/binary scale	literature review and consultation with experts in dementia caregiving. The instrument under	ADL care tasks (α=0.85)
		l	caregiving career	(0=No, 1=Yes)	development to measure unmet need was administered to three groups of dementia CGs based on the	IADL care tasks (α=0.86)
United States		l l	Seven domains:	(Respondents are asked:	"stage" of the care recipient: still living in the community, institutionalized, or deceased. Although	Dementia symptoms (Pearson's
		l	(1) ADL care tasks; (2) IADL		authors do not study the <u>underlying dimensional structure</u> of the PUN measure and do not provide	correlation, r=0.54, p<0.01) (Only two
		l	care tasks; (3) Dementia	with/help providing?"	analyses establishing the <u>concurrent validity</u> of the seven domains or the full scale using simple	items)
		l	symptoms; (4) Timing of	The "yes" responses for	correlations, they conducted three independent multivariate regression path analyses by the "stage" of	Timing of care (α=0.79)
		ŀ	care (5) Formal support	each domain are summed	care recipient to study the associations between unmet needs domains and measures of subjective	Formal support (α=0.77)
			(6) Information; (7)	to create "unmet need"	stress of CGs while controlling for demographic variables. (Three outcome measures of subjective stress	Information (α=0.68)
		l	Confidante/family support	scores.)	were simultaneously examined in each path model: (a) a three-item role overload scale, (b) a three-item	Confidante (α=0.79)
		l		-	role captivity scale, and (c) three-item scale assessing CGs' loss of intimate exchange (feelings of	
		l			emotional/physical separation). All models produced acceptable fit indexes (e.g., RMSEA ranged from	
		l			0.02-0.03 and the GFI ranged from 0.92 to 0.97).	
		l			Among CGs of individuals in the community, scores on the Confidant/family support domain were	
		l			significantly associated with scores on all three outcomes (role overload, role captivity, and loss of	
		l			intimate exchange). For CGs with institutionalized care recipients, scores on the ADL care tasks domain	
		l			were significantly associated with all three outcomes. For those in the deceased care receiver group	
		l			("bereave CGs"), scores on the <i>Confident/family support domain</i> were associated with both <i>role</i>	
		l			overload and loss of intimate exchange.	
Tarlow <i>et al.</i>	ADRD	The Positive	Positive Aspects of	9 items,		Cronbach's α, full scale =0.89.
(2004) <sup>43</sup>			Caregiving	5-point Likert scale	included a measure for positive aspects of caregiving. The studies provided operational definitions of	crombach s a, ran scare c.cs.
(2004)		Caregiving (PAC)		(1=Disagree a lot,		Cronbach's α by subscales:
United States				, •		Self-Affirmation ( $\alpha$ =0.86)
Officed States				3=Neither agree or		Outlook on Life ( $\alpha$ =0.80)
				<u> </u>	accommodate different response options, and (3) instructions were modified to facilitate ease of	Outlook on Life (a=0.80)
					administration. The initial PAC tool contained 11 items.	
			sense that their caregiving	5=Agree a lot)	To establish the <u>structural validity</u> of the 11-item scale, the sample (N=1229) was randomly split into	
			experience is generally		two subsamples to perform a PCA (N=922) and a CFA (N=307). The PCA with oblique rotation and	
		ı	satisfying and rewarding.)		weighted least squares estimation yielded a two-component solution. After eliminating two items with	
		l			low loading the final 9-item scale accounted for 45% of the total variance in items. The CFA indicated a	
		l			good fit for the two-factor solution with a RMSEA statistic of 0.0689.	
		l			Concurrent validity was examined by Spearman's rank correlations between scores in the PAC scale and	
		l			scores in (a) the Somatic and Well-Being subscales of the CES-D, (b) the RMBPC (burden), and (c) the	
		l			Satisfaction with Received Support and Negative Interactions subscales of the Inventory of Socially	
		l			Supportive Behaviors (ISSB). The resulting correlations were significant (p-values < 0.001) and lower	
		l			than expected (all < 0.30, "small to moderate") but in the anticipated directions. The PAC was positively	
		l			associated with wellbeing (rho=0.24) and satisfaction with support (rho=0.15), but negatively associated	
		l			with the RMBPC-burden (rho= -0.23), and somatic aspects of depression (rho= -0.17). PAC was not	
200 1 1					associated with negative social interactions (rho= -0.05, ns).	1:1:1:: (1
Mitrani et al.			, ,	40 items,		Interrater reliability (degree of
(2005)44		· - 1	Two second- or higher-order		, ,	agreement between different raters
			factors:	(ranging from 1=least	Content validity was demonstrated by experienced raters reviewing the coding manual, rating five tapes	- · · · · · · · · · · · · · · · · · · ·
United States		• ,	(1) Intimacy-conflict	adaptive family		calculated with the ICC using the
		,	resolution	<u> </u>	, , , , , , , , , , , , , , , , , , , ,	results from the 46-item first-order
			(2) Freedom from negativity		· · · · · · · · · · · · · · · · · · ·	CFA model with eight factors. ICCs
			Six first-order factors:	functioning)	, , , , , , , , , , , , , , , , , , ,	ranged from 0.617 to 0.937.
			(Intimacy-Conflict		were eliminated resulting in a 60-item scale. A scree plot confirmed a nine-component structure.	
			Resolution)		Iterative CFAs testing alternative models further eliminated items resulting in a 46-item first-order CFA	
			(1) Enmeshment		with eight factors. Subsequent analyses and item deletions yielded a final hierarchical confirmatory	
			(2) Care recipient		factor model with two "second order" factors (labeled as "Intimacy-conflict resolution" and "Freedom	
			disengagement		from negativity") and six first-order factors underlying a 40-item SFSR-DC scale. This hierarchical factor	
			(3) Conflict resolution		model yielded the best fit among competing models (e.g., CFI=0.981, RMSEA=0.048).	
			(4) Expressed positive affect		Concurrent validity was demonstrated by significant ( $p$ -values < 0.001) negative Spearman's correlations	
			(Freedom from Negativity)		between the SFSR-DC "Intimacy-conflict resolution" second-order factor and (a) depression (rho= -0.30)	
			(5) Identified Patienthood		and (b) anxiety (rho= -0.41). "Freedom from negativity" second-order factor was negatively and	

			(6) Expressed anger		significantly associated with subjective burden (rho= -0.30). Depression was measured by the CES-D,	
					anxiety was measured by the State Anxiety Inventory, and subjective burden was measured by RMBPC.	
Gitlin <i>et al.</i>	ADRD	Caregiver	CG reaction to physical	15 items,	The CAFU scale was developed by combining items from two existing scales: eight items from Lawton	Cronbach's α by subscales:
(2005)45		Assessment of	dependence	(Items were scored using	and Brody's (1969)46 instrumental ADL scale and seven items from Hamilton and Fuhrer's (1987)47	ADL dependence scoring (α=0.91)
		Function and	Two factors:	two ordinal scales:	Functional Independence Measure scale. (The CAFU scale was developed to measure both the dementia	ADL upset scoring (α=0.83)
United States		Upset (CAFU)	(1) Activities of Daily Living	Dependence Scale and	patient's level of physical dependence (functional needs) and the CG's reaction (emotional upset) to	ADL mean upset scoring per
			(ADL) dependence and	Upset Scale)	providing assistance with daily activities.)	dependence (α=0.90)
			upset; (2) Instrumental	Dependence scoring scale	To assess the <u>structural validity</u> of the 15-item scale, the sample (N=640) was randomly split into two	. , , ,
			Activities of Daily Living			IADL dependence scoring (α=0.81)
			, ,		component solution explaining 54.7% of the variance. A scree plot confirmed the two components. CFA	, , ,
			upset		with the second subsample further established that the two-factor model was the best fitting model for	
				independence to		dependence (α=0.84)
				•	square error of approximation, RMSEA=0.04).	,
					Concurrent validity was established by significant ( <i>p-values</i> < 0.001) Spearman's rank correlations	
				Upset scoring scale	between CAFU scores and selected criterion measures. CAFU scores (using the Dependence scoring	
					scale) were associated with both vigilance items: more hours on duty (rho=0.24) and more hours doing	
					things for patients (rho=0.24). Greater CG "upset" (using the Upset scoring scale) was significantly	
					correlated with a) more depression (rho=0.32) as measured by the CES-D scale and b) more problem	
				- · · · · · · · · · · · · · · · · · · ·	behavior (rho=0.47), as measured by the RMBPC. Greater CG "upset" was also significantly associated	
				_	with more hours of vigilance for the ADL activities subscale/factor (rho=0.43), but not for the IADL	
				using a 3-point scale.	activities factor.	
Kuhn <i>et al.</i>	Mixed	Knowledge	Knowledge of memory loss,	15 items/questions,	Content validity. A preliminary survey of CGs of individuals in the primary stages of Alzheimer's disease	Cronbach's α, full scale =0.76.
(2005)50		about Memory	Alzheimer's, and related	Each item has 5-response	helped identify three key knowledge domains about memory loss and related care: medical	Cronbach's α by subscales:
		Loss and Care	care	options with a single-	information, caregiving, and legal/financial planning. These domains guided the writing of 31 multiple-	Medical (α=0.46)
United States		test (KAML-C)	Three subscales:	correct answer. Example:	choice items by a panel of seven health professionals.	Caregiving (α=0.61)
			(1) Medical; (2) Caregiving;	Which of the following is	The 31-item pool was administered to three different samples (family CGs, N=45); experts, N=37, and	Legal and financial planning (α=0.53)
			(3) Legal and financial		medical students, N=39). (The sample of medical students was included as a comparison to the experts	
			planning		and the CGs.) Item discrimination and difficulty indexes were calculated using the sample of experts and	Note: The full scale, but not the
				over age 65?		subscales, showed a level of internal
				1. Alzheimer's disease*	, , , , , , , , , , , , , , , , , , , ,	consistency considered acceptable,
					, , , , , , , , , , , , , , , , , , , ,	with a Cronbach's α value above 0.70.
				2. Senility	and related care issues among carers.	
					Group discriminant validity was established by demonstrating the KAML-C's test ability to distinguish	
					between three groups: CGs, experts, and medical students. A Kruskal-Wallis test revealed significant	
				_	differences between the scores of the three groups (p<0.0005) and in post-hoc tests groups scored in	
				5. Benign senescent	the expected order. Experts scored significantly higher than the other two groups (p<0.05), and medical	
				forgetfulness	students scored significantly higher than CGs (p<0.05).	
Gitlin <i>et al.</i>	ADRD	Perceived	State of wellbeing (CG	13 items,		Cronbach's α, full scale =0.90. (Using
(2006) <sup>51</sup>	ADIND		appraisals of self-	5-point Likert scale		half of the sample, N=127)
(2000)			' '	•	potentially, decline, as a consequence of caregiving, which could affect health. A 13-item pool was then	ilan of the sample, N=127)
United States		l' '	· •	I <sup>E</sup>	r · · · · · · · · · · · · · · · · · · ·	Cronbach's α by subscales:
Officed States			Three factors:			Emotional wellbeing (α=0.87)
				•		Physical wellbeing (α=0.79)
					, , ,	, , , , ,
			, , ,	· •	Using the second half of the sample (N=128), concurrent validity was established by significant (p-values	, , ,
			Ability to manage caregiving		< 0.001) Pearson's correlations between PCI scores and a) the CES-D (r=-0.48), b) the Positive Aspects of	
					Caregiving scale scores (r=0.41), and c) the Social Activities Index (r=0.43).	
					<u>Discriminant validity</u> was shown by expected non-statistically significant Pearson's correlations of PCI	
					scores with characterizations of the patients' dementia using the MMSE scores (r=0.01, ns) and activities	
Poilly of al	V D B D	Dartner Dationt	Charad activities between	17 itams (activities)	of daily livingfunctional independence (r=0.07, ns).	Cranhach's a actimates were high for
Reilly <i>et al.</i>	ADRD			17 items (activities),	· · · · · · · · · · · · · · · · · · ·	Cronbach's α estimates were high for
(2006) <sup>52</sup>				5-point Likert scale	anticipatory grief, marital relations, and emotion constructs as well as consultation with an Alzheimer's	
			· · · · · · · · · · · · · · · · · · ·		· · · · · · · · · · · · · · · · · · ·	spouses (0.96).
United States			interference		non-spouse CGs were asked to add activities and judge the frequency, importance, and interference in	
		(PPQSA)			shared activities due to the patient's mood or mental state. Added activities did not differ conceptually	
				ľ	from the originals, so the final PPQSA contained the same original 17 items, yet respondents' input did	
				state interfered with the		
				activity	The PPQSA <u>structural validity</u> was examined through a PCA with Varimax rotation. Authors split the	

				Note: Average PPQSA	sample into spouses (N=71) and non-spouses (N=29) and conducted separate PCA's in each group.	
				score is used as the	Results were similar from both groups yielding one component/factor labeled as relationship	
				"scoring method." CGs	interference.	
				also rate the importance	Some evidence in support of <u>concurrent validity</u> was provided by fitting a multiple regression model	
				of the 17 activities and	using PPQSA scale interference scores as the outcome measure and several criterion scores as	
				the <i>frequency</i> (# of	explanatory variables while controlling for age, gender, and relationship to the patient. The following	
				activities that occurred in	explanatory (criterion) variables were significant predictors (p-values < 0.001) of PPQSA scores:	
				the past 24 hours or the	Caregiver Reaction Assessment, CRA, Work Productivity and Activity Impairment, and Time Spent	
				past week.	Caregiving. All CRA domain scores were also significant predictors of the PPQSA score ( $p$ -values $\leq$ 0.02).	
Menne et al.	ADRD	Decision-Making	Involvement in daily	15 item,	Content validity. Although content validity is not addressed in the current study, prior work is cited <sup>59</sup> on	Cronbach's α, full scale =0.92.
(2008)58		_	decision making	4-point Likert scale	the underlying theories used for DMIS scale development and item adaptation to individuals with	·
(			•	(0=Not at all involved,	dementia and their family CGs.	
United States		' '	decision making	1=A little involved,	The structural validity of the 15-item DMIS scale was established by EFA with a PAF extraction method	
			U	,	and Promax rotation. EFA yielded a unidimensional (one-factor) structure explaining 46.72% of variance.	
			perception of the day-to-day	1	Concurrent validity was demonstrated by expected associations, calculated with Pearson's correlation	
			patient's decision making		coefficients, between total DMIS scores and a) depression, as measured by the CES-D ( $r = -0.16$ , $p < 0.05$ ),	
			involvement.)		b) quality of life, as measured by the Quality of Life-Alzheimer Disease scale (r=0.187, p<0.01), and c)	
			involvement.)		relationship strain, measured by the Dyadic Relationship Scale (r=-0.221, p<0.01).	
Wilks	ADRD	Shortened	Resilience	15 items,		Cronbach's α, full scale =0.89.
(2008) <sup>60</sup>	טווטא		One factor: Global resilience		sample of community-dwelling older adults. The current study examines the psychometric properties of	·
(2006)		(RS-15)	one iactor, giobal resilience	· ·	a shortened 15-item version in a dementia CGs sample.	
United States		(K2-T2)		to 7=Agree)	Structural and concurrent/convergent validity studies were conducted in two separate samples.	
Officed States				to 7-Agree)	· · · · · · · · · · · · · · · · · · ·	
					Structural validity was established through EFA with PAF extraction that yielded a single resilience factor	
					with an eigenvalue of 9.61 and explained 64% of the variance in items.	
					Concurrent validity was demonstrated by significant (p-values < 0.01) Pearson's correlations between	
					scores in the RS-15 scale and scores in the Perceived Stress Scale-10 (r= -0.60) as well as significant	
					correlations with scores in the Perceived Social Support Family Scale (r=0.30) and Perceived Social	
					Support Friends Scale (r=0.34).	
Wilks				10 items,	The PSSS Family and Friends independent "subscales", originally developed by Procidano & Heller	<u>Cronbach's α estimate</u> for Family scale
(2009) <sup>62</sup>			provided by <u>family</u>	5 point Likert scale		was 0.89.
		Support Scale (S-		(ranging from 0=Strongly		Cronbach's α by subscales (Family
United States			(1) Relationship,	disagree to 4=Strongly	· ·	scale):
			Togetherness; (2) Moral,	agree)		Relationship, Togetherness (α=0.82)
			emotional support; (3)		Alzheimer's CGs was randomly split into two samples. The first half (N=115) was administered the	Moral, emotional support (α=0.79)
		(Scale appraising	Openness, reliance			Openness, reliance (α=0.79)
		family support)				Guttman's split-half reliability estimate
					the same underlying three-factor structure and similar patterns of factor loadings across factors. The	for the Family scale was 0.92.
					proportion of variance explained was 74% for the Family scale.	
					Concurrent validity was demonstrated by significant negative Pearson's correlations between scores in	
					the S-PSSS "Family" scale and scores in the Perceived Stress Scale (r= -0.18, p<0.05) as well as significant	
					positive correlations with scores in the Resilience Scale (r=0.15, p<0.05).	
		Shortened	Perceived social support as	10 items,	Structural validity. The EFA with the PAF extraction method and Varimax rotation also yielded a three-	Cronbach's α estimate, Friends scale
		Perceived Social	provided by <u>friends</u>	5 point Likert scale	factor structure explaining 76% of the variance in items for the "Friends" scale.	=0.90.
		Support Scale (S-	· · · · · · · · · · · · · · · · · · ·	·	Concurrent validity was demonstrated by significant negative correlations between scores in the S-PSSS	Cronbach's α by subscales (Friends
			(1) Relationship,		Friends scale and the Perceived Stress Scale (r= -0.26, p<0.05) as well as significant positive correlations	
			• •	agree)		Relationship, Togetherness (α=0.86)
			emotional support; (3)	"	i i i i i i i i i i i i i i i i i i i	Moral, emotional support (α=0.79)
			Openness, reliance			Openness, reliance (α=0.81)
		friends support)	1			Guttman's split-half reliability, Friends
						scale = 0.94.
Carpenter <i>et</i>	Mixed	The Alzheimer's	Knowledge of Alzheimer's	30 items,	Content validity. The Alzheimer's disease Knowledge Scale (ADKS) is an update to the 30-year-old	Cronbach's α, full scale =0.71
al. (2009) <sup>65</sup>			disease	2-point, binary scale		Test-retest reliability for a subsample
an. (2003)		Knowledge Scale		(0=False, 1=True)		(N=40) at an interval of 2 to 50 hours
United States			(1) Risk factors; (2)	(0-1 alse, 1-11ue)		between tests (r=0.81, p<0.001).
Officed States						Guttman's split-half reliability estimate
			Assessment and diagnosis;			for the full scale=0.55 (p<0.001).
			(3) Symptoms; (4) Course; (5) Life impact; (6)		Before studying the psychometric properties of the full scale, authors first analyzed individual item properties via item discrimination indexes, item difficulty indexes, and item homogeneity using split	ioi the full scale=0.55 (p<0.001).
		1	ic) rite impact; (p)	İ	properties via item discrimination indexes, item difficulty indexes, and item nomogeneity using split	

_		 				
()	Czaja <i>et al.</i> 2009) <sup>67</sup> United States Gickrey <i>et al.</i> 2009) <sup>70</sup> United States	REACH Risk Appraisal Measure (RAM)  Caregiver- targeted quality- of-life (CGQOL)	Caregiving; (7) Treatment and management  CG risk Six domains: (1) Depression; (2) Burden (3) Self-care and healthy behaviors; (4) Social support (5) Safety; (6) Patient problem behaviors  CG Quality-of-Life Three higher order factors Ten domains/subscales: Tangible Assistance (1) Assistance in ADLs (2) Assistance in IADLs (3) Personal time (4) Role limitations due to caregiving Psychosocial (5) Family involvement (6) Caregiving demands (7) Worry (8) CG feelings Benefits/Faith (9) Spirituality and faith (10) Benefits of caregiving	16 items, (Mixed scale formats) 2-point/binary scale (0=No, 1=Yes), 3-point Likert scale (0=Never to 2=Often), 4-point Likert scale (from 0=Not at all to 3=Very), 5-point Likert scale (from 0=Poor to 4=Excellent), 6-point Likert scale (from 0=Never to 5=Nearly always) 80 items, Items have different scales and response categories. Note: The 80 items are distributed across 10 subscales. The final scoring for the CGQOL scale recodes the initial response categories into a 0-100 rating where	The <u>structural validity</u> of the CGQOL was established by iterative EFAs with Promax rotations. Guttman's weakest lower bound, Cattell's scree plot, and parallel analysis were examined to determine the number of factors. A final <u>higher order factor analysis</u> identified a three-factor solution influencing the 10-subscales or factors. The three higher order factors were interpreted as: <i>Tangible assistance</i> , <i>Psychosocial</i> , and <i>Benefits/faith</i> . Associations between the three factors ranged from 0.04 to 0.52. Multitrait scaling was used to examine item and subscale internal consistency estimates, item-scale correlations, and correlations among scales. This process reduced the scale from 91 to 80 items. Multitrait-multimethod analysis was used to assess the validity of the scale by examining the correlations between multiple traits measured using the 10 subscales.  Concurrent validity was demonstrated by significant negative correlations between hours-per-week caregiving and all subscales (r=0.14 to 0.68; <i>p-values</i> < 0.01) except Caregiving Benefits and Spirituality/Faith (0.092, <i>p</i> >0.05). The association of duration of being a CG and IADLs was also significant (= -0.192, <i>p</i> =0.007; r= -0.163, <i>p</i> =0.02).	Cronbach's $\alpha$ , full scale =0.65.  Cronbach's $\alpha$ by subscales: Assistance in ADLs ( $\alpha$ =0.88); IADLs ( $\alpha$ =0.93); Personal Time ( $\alpha$ =0.78); Role Limitations ( $\alpha$ =0.86); Caregiving Demands ( $\alpha$ =0.86); Worry ( $\alpha$ =0.82); CG Feelings ( $\alpha$ =0.94); Spirituality/Faith ( $\alpha$ =0.92); Benefits of caregiving ( $\alpha$ =0.89). Test-retest reliability (within 21 days) was calculated with the ICC with N=38. Test-retest reliability by subscales: Assistance in ADLs (ICC=0.86); IADLs (ICC=0.86); Personal Time (ICC=0.63); Role Limitations (ICC=0.53); Family Involvement (ICC=0.74); Caregiving Demands (ICC=0.72); Worry (ICC=0.53); G Feelings (ICC=0.65);
			(7) Worry (8) CG feelings <u>Benefits/Faith</u> (9) Spirituality and faith		caregiving and all subscales (r=0.14 to 0.68; $p$ -values < 0.01) except Caregiving Benefits and Spirituality/Faith (0.092, $p$ >0.05). The association of duration of being a CG and IADLs was also significant (= -0.192, $p$ =0.007; r= -0.163, $p$ =0.02).	Assistance in ADLs (ICC=0.86); IADLs (ICC=0.86); Personal Time (ICC=0.63); Role Limitations (ICC=0.53); Family Involvement (ICC=0.74); Caregiving Demands (ICC=0.72); Worry
( (	pstein- ubow <i>et al.</i> 2010) <sup>71</sup> United States	Assessment Questionnaire (CSAQ)		(Mixed item scales) 2-point/binary scale (0=No, 1=Yes); 10-point Likert scale (ranging from 1=Not Stressful to 10=Extremely Stressful) or (from 1=Very healthy to 10=Very ill)	Note: The CSAQ was originally developed and tested by the American Medical Association (AMA) targeting a general population of family CGs. AMA reported a Cronbach's $\alpha$ reliability of 0.78 during scale development. To our knowledge, no further details on content validation and underlying factorial structure have been reported. The field study by Epstein and Lubow (2010) reported here examined the concurrent and predictive validity of the CSAQ scale in a sample of 106 predominantly (91.5%) dementia CGs. Assuming unidimensionality, a "total" score for the CSAQ was used to report the results. The concurrent validity of the CSAQ was demonstrated by a significant positive Pearson's correlation with the CES-D (r=0.807, $p$ <0.001). Similar significant positive associations (all $p$ <0.001) were found between CSAQ and a) stress measured by the Rapid Screen for Caregiver Burden (r=0.707), b) grief, measured with the Inventory for Traumatic Grief, Pre-Loss Version (r=0.594), and c) stress assessed with the Perceived Stress Scale-4-Item Version (r=0.682). CSAQ's scores sensitivity to predict significant depressive symptoms was 0.98, with a specificity= 0.52.	Cronbach's α, full scale =0.82.
	ough <i>et al.</i> 2010) <sup>72</sup>		Spirituality One factor: Intrinsic	6 items, 11-point scale (ranging	Content validity of the scale was established through literature reviews on caregiving burden and spirituality as a coping resource. Authors use the ISS scale, originally developed by Hodge, (2003) <sup>73</sup> , and	Cronbach's α, full scale =0.919.
			<u> </u>			

Position of States   Position States   Positio			r		1		
Security of the company of the com			(ISS)	spirituality	from 0=Spirituality	i i i	
Dis-particular salidate was established by EPF with PAP exhausts or evaluating in a single-dimension solution behaviorable provided by significant composition of passing a frequency of view for Paper and a Measurement of the Paper (CAP) and AB provided and a manufacture of the Paper (CAP) in the content of the Paper (CAP) in the Caregiver's Assessment of the Caregiver's Assessment of the Paper (CAP) in the Caregiver's Assessment of the Caregiv	United States	5			·		for the full scale=0.914.
bibout life) submit life with the bibout life with bibout life bib							
bout lifty							
and all frequency of parset (nd.55), b) the Private Prayer as a Means of Coping (UPPMC, mc.65), b), bits place in the Private Prayer as a Means of Coping (UPPMC, mc.65), bits place in the Private Prayer as a Means of Coping (UPPMC, mc.65), bits place in the Private Prayer as a Means of Coping strong Froblet Prayer and Strong Froblet Prayer and Strong Fr						1	
Private Progres as A Means of Coping scores (mGS), and of the Common-discolor Resilients Scale scores (mg GS) and proceedings of \$5 copies with the \$8 stores were not exocuted with relation to corn recipient (mGS), p=0.43, p=0.31, p=0.3					about life)	· · · · · · · · · · · · · · · · · · ·	
Feder et al.   ADRD   Caregiver   Caregivers   Caregiver   Caregivers   Caregiver							
Secundariany ADD  Nontgemeny and Collection  Secundariany ADD  Nontgemeny and Collection  Secundariany ADD  Nontgemeny and Collection  Interest factors  Completed  Congrigion  MB-CBS  NB-CBS							
Source of the Patient (CAP)  Lined States  Drilled States  Dri							
Support of Caregore   Caregore							
Subjective burden, (2)   Expression   1 (Disjective burden, (2)   Expression   2 (Disjective stress burden   2 (Disjective stress   2 (Disje		ADRD	,				•
Burden Sole   Subjective demand or relationship burden (SS scale).   MBC 653   value demand or sole relationship burden (SS scale).   MBC 653   value demand or sole relationship burden (SS scale).   MBC 653   value demand or sole relationship burden (SS scale).   MBC 653   value demand or sole relationship burden (SS scale).   MBC 653   value demand or sole relationship burden (SS scale).   MBC 653   value demand or sole relationship burden (SS scale).   MBC 653   value demand or sole relationship burden (SS scale).   MBC 653   value demand or sole relationship burden (SS scale).   MBC 653   value demand or sole relationship burden, and objective burden (confirming the same factor structure and that the association between the subscales and solid burden (SS scale).   MBC 653   value demand or sole relationship burden, and objective burden (confirming the same factor structure and that the association between the subscales and solid burden (scale).   MBC 653   value scales of carginging spots and solid burden (scales).   Value bu			. 0				
Mis-CSS    selationship burden; (3)   Subjective stress burden   Subjecti	(2011)/8		_	I	·		, ,
Subjective stress burden Results revealed that the MB-CSE factor structure had configurated and metric invariance across the samples of caregiveries in the measurement of stress burden (are 1931) headers and shall the saccisation between headers and shall thilliers. That is, the interpretation of scale items can be considered consistent across these two groups of caregiver per shall be the same across spouse and adult withing. That is, the interpretation of scale items can be considered consistent across these two groups of caregivers. The results showed that the MB-CSE depictive burden considered consistent across these two groups of caregivers. The results showed that the MB-CSE depictive burden subsidies and ARIS were significantly associated. Problem behavior scores were also fall building to the patient of sealth.  Ender et al. (ADED Caregivers' Perceived Burden COP)  Dinied States  United States  United States  United States  United States  United States  United States  United States  CPROI;  Caregivers' Assessment of Control of the patient (CAP)  United States (CAP)  United Sta				,			` ,
samples of caregiving spouses and adult children in the measurement of stress burden, relationship burden, and objective burden confirming the same factor structure and that the association between back item and the latent construct it measures per subscale is the same across spouses and adult. Relational burden (acro.83) higher than the construction of caregiver.  Nate: To provide some evidence of "criterion validity", authors test hypothesized relationships between the subscales and known caregiving burden measures fitting two separate structural equations models. The results showed that the MPE GS objective burden subscales and ADIS were significantly associated. Problem behavior scores were also significantly associated with all three MPE GS objective burden factors, and a subscales and known caregiving burden measures fitting two separate structural equations models. The results showed that the MPE GS objective burden subscales and ADIS were significantly associated. Problem behavior scores were also significantly associated with all three MPE GS objective burden subscales. Burden factors is not always with the soque and failines and secure for results. Burden factors is not always with the soque and failines and secure for results. Burden factors is not always with the soque and failines, and secure for remaining failines, and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines, and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for remaining failines and secure for	United States	5			5=A lot more)		· · · · · · · · · · · · · · · · · · ·
burden, and objective burden confirming the same factor structure and that the association between heat them and the latest construct it measures pre subscale is the same across spouse and adult whilefren. That is, the interpretation of scale items can be considered consistent across these two groups of carens.  Note: To provide some evidence of "criterion validity", authors test hypothesized relationships between the subscales and known caregiving burden subscales and ADLs were significantly associated. Problem behavior scores were also significantly associated with a little were significantly associated. Problem behavior scores were also significantly associated with a little were significantly associated. Problem behavior scores were also signified the same pattern of results.  Ender et al. ADRD  Caregivers' Assessment of Burden (2AP)  United States  United States  United States  CPRQL: CPRQL: CPRQL: CPRQL: CPRQL STATE AND STATE				Subjective stress burden		, ,	
ash item and the latent construct it measures per subscale is the same across spouses and adult, thidren. That is, the interpretation of the interpretation construct of the subscale is the same across spouses and adult, thidren. That is, the interpretation of the interpretation of the subscale construction of the subscale subscale is the same across spouses and adult, which is the subscale subscale in the subscale subscale is the same across sheet was good prices burden (ac-0.90) soft cares;  Note: To provide some evidence of "criterion validity", authors test hypothesized relationships between the subscales and known caregiving burden measures fitting two separate structural equations models. The results showed that the MEACS burden factors, Both analyses with the pages of the Perceived of the Patient (CAP)  United States  One that the provided of the page of the Patient (CAP)  United States  One that the page of the page of the page of the Patient (CAP)  United States  One that the page of the pag							· · · · · · · · · · · · · · · · · · ·
hildren. That is, the interpretation of scale Items can be considered consistent across these two groups of carers.  Neter To provide some evidence of "criterion validity", authors test hypothesized relationships between the substrales and known caregiving buffer measures fifting two separate structural equations models. The results showed that the Mis-CBS-objective burden substrale and ADIs were significantly associated with a structural equations models. The results showed that the Mis-CBS-objective burden substrale and ADIs were significantly associated with all three Mis-CBS buffer factors. Both provided the same pattern of results.  Concent validity. The assessment goals and measurement channels of the CPBQ were initially informed in the pattern of results. Concent validity in the pattern (CAP) Exercised Exercised Continuing. The initially analyzed CPBQ items of the pattern (CAP) Exercised Communication, Social ability of the pattern (CAP) CPBQ. Scale 2: Interest analysis (not labeled) Exercised Concent validity. The pattern (CAP) CPBQ. Scale 2: Interest analysis (not labeled) Exercised Concentration of the pattern (CAP) Construct.)  Assessment of the Patient (CAP) Construct.)  Assessment of the Patient (CAP) Construct.)  Assessment of the patient (CAP) Construct.)  Assessment of the patient (CAP) Construct.)  Assessment of the patient (CAP) Construct.)  Assessment of the patient (CAP) Construct.)  Assessment of the patient (CAP) Construct.)  Assessment of the patient (CAP) Construct.)  Assessment of the patient (CAP) Construct.)  Assessment of the patient (CAP) Construct.)  Assessment of the patient (CAP) Construct.)  Assessment of the patient's consideration of the patient's consideration of the patient's consideration of the patient's consideration of the patient's consideration of the patient's consideration of the patient's consideration of the patient's consideration of the patient's consideration of the patient's consideration of the patient's consideration of the patient's consideration of the pati							, , ,
Fider et al. ADRO (2012)**  United States Output of States of Caregiver's Assessment of Examples of Caregiver's Assessment of Examples of							, , ,
Note: To provide some evidence of "criterion validity", authors test hypothesized relationships between the subscales and ADIs were significantly associated problem behavior screen were larged that the MB-CBS objective burden subscale and ADIs were significantly associated problem behavior screen were larged with all three MB-CBS burden factors. Both analyses with the spouses and children samples yielded the same pattern of results.  Likert scale (cut-points or behavior screen were larged to the pattern of results.  Likert scale (cut-points or behavior screen were larged to the same pattern of results.  Likert scale (cut-points or behavior screen were larged to the same pattern of results.  Likert scale (cut-points or behavior screen were as follows: Caregiver-perceived patient thresholds not provided burden analysis of the same pattern of results.  Likert scale (cut-points or behavior screen and the same pattern of results.  Likert scale (cut-points or behavior screen and the same pattern of results.  Likert scale (cut-points or behavior screen and scales). The same scales and measurement of the same pattern of results.  Likert scale (cut-points or behavior screen and scales). The same scales of such analysis scale of united review of the results and the tem content analysis, the seasons of the same pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.  Likert scale (cut-points or the pattern of results.)  Likert scale						<u> </u>	Stress burden (α=0.90)
The subscales and known caregiving burden measures fitting two separates structural equations models. The results showed that the Mac Schopletive burden subscale and ADS were significantly associated. Problem behavior scores were also significantly associated with all three MB-CBS burden factors. Soth analyses with the spouses and fire sample syleded the same pattern of results.    Caregiver- perceived Burden   Questionnaire (CPBQ)   Caregiver-perceived patient, thresholds not provided functional engagement)   Questionnaire (CPBQ)   Caregiver-perceived patient, for the Patient (CAP)   Caregiver-perceived p							
The results showed that the MB-CBS objective burden subscale and ADLs were significantly associated. Problem behavior scores were observed to specify the MB-CBS burden (2012)**  Erder et al. ADRD (aregiver-perceived patient (CAP) (Caregiver-perceived patient thresholds not provided) to be patient (CAP) (Caregiver-perceived patient thresholds not provided) (Caregiver-perceived patient thresholds not provided) (Caregiver-perceived patient thresholds not provided) (PA) (Caregiver-perceived patient thresholds not provided) (PA) (Caregiver-perceived patient thresholds not provided) (PA) (Caregiver-perceived patient thresholds not provided) (PA) (Caregiver-perceived patient) (PA) (Caregiver-perceived patient) (PA) (Caregiver-perceived patient) (PA) (Caregiver-perceived patient) (PA) (Caregiver-perceived patient) (PA) (PA) (Caregiver-perceived patient) (PA) (PA) (Caregiver-perceived patient) (PA) (PA) (Caregiver-perceived patient) (PA) (PA) (Caregiver-perceived patient) (PA) (PA) (Caregiver-perceived patient) (PA) (PA) (PA) (PA) (PA) (PA) (PA) (PA							
Problem behavior scores were also significantly associated with all three MB-CBS burden factors. Both analyses with the spouses and children samples yielded the same parter or fesults.  Erder et al. ADRD Caregiver-Perceived Burden (2012)**  Dritted States (2012)**  Unit							
Series et al.   ADRD   Caregiver-   Caregi							
Caregiver   Caregiver   Caregiver   Caregiver   Caregiver   Assessment of the Patient (CAP)   Liters tasel (cut-points or in the patient (CAP)   Liters tasel (cut-points or in the patient (CAP)   Liters tasel (cut-points or in the patient (CAP)   Liters (CAP)						į ,	
Perceived Burden   Dried States   Perceived Burden   Caregiver-perceived patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-from the   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceive patient thresholds not provided   Caregiver-perceived burden   Caregiver-							
United States    Burden   Questionnalize (CPBQ):   Exhausting the procession of 2 purposes of the patient (CPBQ):   Exhausting the procession of 2 purposes of the patient (CAP) scale 1:   Exhausting the patient (CAP) scale 1:   Exhausting the patient (CAP) scale 1:   Exhausting the patient (CAP) scale 1:   Exhausting the patient (CAP) scale 3 purposes of the patient (CAP) scale 3 purposes of the patient (CAP) scale and a 13-tem Caregivers' Assessment of the patient (		ADRD		"	· · · · · · · · · · · · · · · · · · ·	· · · · · · · · · · · · · · · · · · ·	
United States  Questionnaire (CPBQ): Three "factors" from the Scale 1: FFA analysis (not labeled) Caregivers' Assessment of the Patient (CAP) sought of the Patient (CAP)	(2012)81			' '			<u>.</u>
CPBQ]:					thresholds not provided)		
Scale 1: Caregiver' Assessment of unidimensional (one-factor) the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Patient (CAP) Exercised and a season and the Season and the	United States	6				ļ.	
Caregivers' Assessment of the Patient (CAP) construct.)  Caregivers' Assessment of the Patient (CAP) construct.)  Caregivers' Caregivers' Assessment of the Patient (CAP) construct.)  Caregivers' Car			, ,				
Assessment of the Patient (CAP) construct.)  Assessment of the Patient (CAP) construct.)  Assessment of the Patient (CAP) construct.)  Assessment of the Patient (CAP) construct.)  Assessment of the Patient (CAP) construct.)  Assessment of the Patient (CAP) construct.)  Assessment of the Patient (CAP) construct.  Assessment of the Patient (CAP) construct.  Assessment of the Patient (CAP) construct.  Assessment of the Patient (CAP) construct.  Assessment of the Patient (CAP) construct.  Assessment of the Patient (CAP) construct.  Assessment of the Patient (CAP) construct.  Assessment of the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Surdent in the Patient (CAP) construct.  Assessment of Perceived Surdent in the Patient (CAP) construct.  Assessment of Perceived Surdent in the Patient (CAP) construct.  Assessment of Perceived Surdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Surdent in the Patient (CAP) construct.  Assessment of Perceived Surdent in the Patient (CAP) construct.  Assessment of Perceived Surdent in the Patient (CAP) construct.  Assessment of Perceived Surdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Surdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) construct.  Assessment of Perceived Durdent in the Patient (CAP) const				, ,		• • •	· · · · · · · · · · · · · · · · · · ·
After deleting items with low loadings, 20-items were retained for CAP. A CFA was executed on the second split-half sample. The model failed tests of comparative fit index (CFI=0.863), root mean square error of approximation (RMSEA=0.073), and standardized root mean square residual (SRMR=0.065), but "items were judged by the experts as the most plausible and meaningful".  Next, asks chanalysis of the CAP scale was conducted showing good overall fit suggesting that it measured a single underlying construct, as the Rasch model assumes unidimensionality.  Concurrent validity was shown by significant Spearman's rank correlations (p-values-c0.001) between the CAP and the NPI (rho=0.38), the Severe Impairment Battery (rho=-0.45), that Altheimer's Disease Cooperative Study-ADL Scale (rho=-0.57), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.45), and the Functional Assessment Staging Tool (rho=0.45).  Caregiver-Perceived Themselves (CAT)  (Caregiver-Service)  Themselves (CAT)  (Caregiver-perceived burden thresholds not provided)  Questionnaire  (CPBQ):  engagement)  Scale 2:  vo "factors" from the EFA  Also, the Rasch analysis of the 10-item CAT scale showed good overall fit suggesting a single  (unidimensional) construct.  Concurrent validity was shown the validative was shown the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, and SRMR=0.056) yet again "items were judged by the experts as the most plausible and meaningful".  Also, the Rasch analysis of the 10-item CAT scale showed good overall fit suggesting a single  (CPBQ):  engagement)  Scale 2:  vo "factors" from the EFA  Also, the Rasch analysis of the 10-item CAT scale showed good overall fit suggesting a single  (unidimensional) construct.  Concurrent validity was demonstrated by significant Spearman's rank correlations (p<0.001) between the CAT and the NPI (rho=0.35), the Severe Impairment Battery (rho=0.14), the Clinician's Interview-Based Impression of Change-Plus Caregiver In			_				
second split-half sample. The model failed tests of comparative fit index (CFI=0.863), root mean square error of approximation (RMSEA=0.073), and standardized root mean square residual (SRMR=0.065), but "litems were judged by the experts as the most plausible and meaningful".  Next, a Rasch analysis of the CAP scale was conducted showing good overall fit suggesting that it measured a single underlying construct, as the Rasch model assumes unidimensionality.  Concurrent validity was shown by significant Spearman's rank correlations (p-values-0.001) between the CAP and the NPI (rho=0.38), the Severe Impairment Battery (rho=-0.45), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=-0.57), the Clinician's Interview-Based Impression of Change-Plus Caregiver-Perceived burdenthresholds not provided) on the second split-half sample yielding a 2-factor [Caregiver-perceived burdenthresholds not provided) on the second split-half sample with low loadings, 10-items were retained for CAT. A CFA was conducted structure. After deleting items with low loadings, 10-items were retained for CAT. A CFA was conducted structure. After deleting items with low loadings, 10-items were retained for CAT. A CFA was conducted with the caregiver-perceived burdenthresholds not provided) in relation to the patient's engagement)  Caregiver-perceived burdenthresholds not provided) in relation to the patient's engagement by segment of the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, and SRMR=0.056) yet again "items were judged by the experts as the most plausible and meaningful".  Also, the Rasch analysis of the 10-item CAT scale showed good overall fit suggesting a single (unidimensional) construct.  Caregivers' Assessment of Themselves unidimensional (lone-factor) construct.  Caregivers' Assessment of Themselves unidimensional (lone-factor) construct.  Concurrent validity was demonstrated by significant Spearman's rank correlations (p-0.01) between the CAT and the NPI (rho=0.35), the Sev				, , ,			
error of approximation (RMSEA=0.073), and standardized root mean square residual (SRMR=0.065), but "items were judged by the experts as the most plausible and meaningful".  Next, a Rasch analysis of the CAP scale was conducted showing good overall fit suggesting that it measured a single underlying construct, as the Rasch model assumes unidimensional (one-factor)  Concurrent validity was shown by significant Spearman's rank correlations (p-values<0.001) between the CAP and the NPI (rho=0.38), the Severe Impairment Battery (rho=-0.45), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=-0.57), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.45), and the Functional Assessment Staging Tool (rho=0.36).  The structural validity for CAT was established by EFA with on a split-half sample, rhose of the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, on relation to the patient's engagement)  Caregiver-perceived burden thresholds not provided) on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA			the Patient (CAP)	construct.)			
"items were judged by the experts as the most plausible and meaningful".  Next, a Rasch analysis of the CAP scale was conducted showing good overall fit suggesting that it measured a single underlying construct, as the Rasch model assumes unidimensionality.  Concurrent validity was shown by significant Spearman's rank correlations (p-values<0.001) between the CAP and the NPI (rho=0.38), the Severe Impairment Battery (rho=0.45), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=0.57), the Clinician's Interview-Based Impression of Change-Plus Caregiver-Perceived Impression of Change-Plus Caregiver Input (rho=0.45), and the Functional Assessment Staging Tool (rho=0.36).  The structural validity for CAT was established by EFA with on a split-half sample yielding a 2-factor Likert scale (cut-points or hersholds not provided) and SRMR=0.056) yet again "items were retained for CAT. A CFA was conducted on thresholds not provided) in relation to the patient's engagement)  Caregiver-perceived burden thresholds not provided)  Questionnaire (CPBQ): engagement)  Scale 2: Two "factors" from the EFA (unidimensional) construct.  Caregiver's Assessment of Themselves (CAT)  Caregiver's Assessment of Internal vas calculated with the CC0-0.58.  Also, the Rasch analysis of the CAP scale was conducted showing good overall fit suggesting a single (unidimensional) construct.  Concurrent validity was shown by significant Spearman's rank correlations (p-0.001) between the Concurrent validity was demonstrated by significant Spearman's rank correlations (p-0.001) between the Concurrent validity was salvery and relevance as a measure of self-efficacy for decision of Change-Plus Caregiver Input (rho=0.23), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.23), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.23), the Severe Impairment Battery (rho=0.14).  Lopez & ADRD Surrogate Self-efficacy for decision Si tems, 4-point Likert scale interval validity was established by							
Next, a Rasch analysis of the CAP scale was conducted showing good overall fit suggesting that it measured a single underlying construct, as the Rasch model assumes unidimensionality.  Concurrent validity was shown by significant Spearman's rank correlations (ρ-αμενε-0.001) between the CAP and the NPI (rho=0.38), the Severe Impairment Battery (rho=-0.45), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=-0.57), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.45), and the Functional Assessment Staging Tool (rho=0.36).  The structural validity for CAT was established by EFA with on a split-half sample yielding a 2-factor Themselves (CAT) Likert scale (cut-points or Structure. After deleting items with low loadings, 10-items were retained for CAT. A CFA was conducted On the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, and SRMR=0.056) yet again "items were judged by the experts as the most plausible and meaningful". Ca-0.58.  (CPBQ): engagement) Also, the Rasch analysis of the 10-item CAT scale showed good overall fit suggesting a single (unidimensional) construct.  Assessment of Themselves (CAT) In relation to the patient's analysis (not labeled) (Rasch analysis of the 10-item CAT scale showed good overall fit suggesting a single (unidimensional) construct.  Concurrent validity was demonstrated by significant Spearman's rank correlations (ρ<0.001) between the CAT and the NPI (rho=0.35), the Severe Impairment Battery (rho=0.01), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=0.24), the Clinician's Interview-Based Impression of Change-Plus (CAT) Construct.)  Lopez & ADRD Surrogate S							
measured a single underlying construct, as the Rasch model assumes unidimensionality.  Concurrent validity was shown by significant Spearman's rank correlations (p-values>c.0.001) between the CAP and the NPI (rho=-0.48), the Severe Impairment Battery (rho=-0.45), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=-0.57), the Clinician's Interview-Based Impression of Change-Plus Caregivers of Themselves (CAT)  Perceived Burden Questionnaire (CPBQ): engagement) Caregivers' Assessment of Interview-Bound of the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, and SRMR=0.056) yet again "items were judged by the experts as the most plausible and meaningful". Caregivers' Assessment of Interview-Bound of the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, and SRMR=0.056) yet again "items were judged by the experts as the most plausible and meaningful". Caregivers' Assessment of Interview-Bound of the Scale 2: Two "factors" from the EFA (unidimensional) construct. Caregivers' Assessment of Interview-Based Impression of Change-Plus Concurrent validity was demonstrated by significant Spearman's rank correlations (p<0.001) between the CAT and the NPI (rho=0.35), the Severe Impairment Battery (rho=-0.19), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=-0.24), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Lopez & ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Surrogate Decision Making  ADRD Su							
Concurrent validity was shown by significant Spearman's rank correlations (p-values<0.001) between the CAP and the NPI (rho=0.38), the Severe Impairment Battery (rho=-0.45), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=-0.57), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.45), and the Functional Assessment Staging Tool (rho=0.36).    Caregiver- Caregivers' Assessment of Perceived Themselves (CAT) (Caregiver-perceived burdent on the patient's Burden (Caregiver-perceived burdent on the patient's engagement) (Caregiver-perceived burdent to the						·	
the CAP and the NPI (rho=0.38), the Severe Impairment Battery (rho=-0.45), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=-0.57), the Clinician's Interview-Based Impression of Change-Plus Caregiver- Perceived Burden Questionnaire (CPBQ): engagement) Scale 2: Two "factors" from the EFA Caregivers' Assessment of Themselves (CAT) (unidimensional) construct. Caregivers' Assessment of Themselves (Likert scale (cut-points or thresholds not provided) Scale 2: Two "factors" from the EFA Analysis (not labeled) (Rasch analysis suggested a Unidimensional (one-factor) (CAT) (CAT)  Construct.)  Likert scale (cut-points or thresholds not provided) Themselves (Likert scale (cut-points or thresholds not provided) The structural validity for CAT was established by EFA with on a split-half sample yielding a 2-factor The structural validity for CAT was established by EFA with on a split-half sample yielding a 2-factor The structural validity for CAT was established by EFA with on a split-half sample yielding a 2-factor The structural validity for CAT was established by EFA with on a split-half sample yielding a 2-factor The structural validity for CAT was established by EFA with on a split-half sample yielding a 2-factor The structural validity for CAT was established by EFA with on a split-half sample yielding a 2-factor The structural validity after a 4-week on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, and SRMR=0.056) yet again "items were judged by the experts as the most plausible and meaningful".  ICC=0.58.  Likert retest reliability after a 4-week on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, and SRMR=0.056) yet again "items were judged by the experts as the most plausible and meaningful".  ICC=0.58.  Caregiver In model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, and SRMR=0.056) yet again "items were judged by the experts as the most plausible and meaningful".  ICC=0.58.  Concurrent vali							
Caregiver- Perceived Burden Questionnaire (CPBQ): engagement) Caregivers' Assessment of Themselves (CAT) Caregivers' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CPBQ): engagement) CAREGIVERS' Assessment of Themselves (CPBQ): engagement) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CPBQ): engagement) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of Themselves (CAT) CAREGIVERS' Assessment of CAT and the NPI (rho=0.35), the Severe Impairment Battery (rho=-0.19), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=-0.24), the Clinician's Interview-Based Impression of Change-Plus (CAT) CAT) CATOR							
Caregiver Input (rho=0.45), and the Functional Assessment Staging Tool (rho=0.36).  Caregiver-Perceived Themselves (CAT) (Caregiver-perceived burden Questionnaire (CPBQ): engagement) (Caregivers' Assessment of Themselves (CAT) (Two "factors" from the EFA analysis (not labeled) (Rasch Assessment of Themselves (CAT) (							
Caregiver- Caregivers' Assessment of Themselves (CAT)  Burden (Caregiver-perceived burden in relation to the patient's engagement)  Caregivers' Assessment of Themselves (CAT)  Caregivers' construct.  Caregivers' Assessment of Themselves (CAT)  Caregiver-perceived burden in relation to the patient's engagement (Caregivers' analysis (not labeled) (Rasch analysis suggested a unidimensional (one-factor) (CAT)  CASSESSMENT OF Themselves (CAT)  Caregivers' Assessment of Themselves (CAT)  Caregivers' Assessment of Themselves (CAT)  Caregivers' Assessment of Themselves (CAT)  Caregivers' Analysis (not labeled) (Rasch analysis suggested a unidimensional (one-factor) (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  COncurrent validity was demonstrated by significant Spearman's rank correlations (p<0.001) between the CAT and the NPI (rho=0.35), the Severe Impairment Battery (rho=0.19), the Alzheimer's Disease (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF Themselves (CAT)  CASSESSMENT OF THEMSELOP (CAT)  CASSESSMENT OF							
Perceived Burden (Caregiver-perceived burden (CPBQ): engagement) (Caregivers' analysis (not labeled) (Rasch Assessment of Themselves (CAT) (CAT					40.1		
Burden Questionnaire (CPBQ): in relation to the patient's engagement)  Two "factors" from the EFA analysis (not labeled) (Rasch analysis suggested a unidimensional (one-factor) (CAT)  Lopez & ADRD Guarino  Burden (Caregiver-perceived burden thresholds not provided) in relation to the patient's engagement thresholds not provided) on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, in relation to the patient's engagement)  Also, the Rasch analysis of the 10-item CAT scale showed good overall fit suggesting a single (unidimensional) construct.  Concurrent validity was demonstrated by significant Spearman's rank correlations (p<0.001) between the CAT and the NPI (rho=0.35), the Severe Impairment Battery (rho=-0.19), the Alzheimer's Disease unidimensional (one-factor) construct.)  Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Lopez & ADRD Surrogate Guarino  Burden (Caregiver-perceived burden thresholds not provided) on the second split-half sample. The model produced a satisfactory fit (e.g., CFI=0.918, RMSEA=0.084, and SRMR=0.056) yet again "items were judged by the experts as the most plausible and meaningful".  Also, the Rasch analysis of the 10-item CAT scale showed good overall fit suggesting a single (unidimensional) construct.  Concurrent validity was demonstrated by significant Spearman's rank correlations (p<0.001) between the Rasch model) estimate for the full scale=0.83.)  Scale 2: Two "factors" from the EFA (unidimensional) construct.  Construct.  Caregiver's construct.  Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Construct.  Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Caregiver Input (rho=0.24), the Clinician's Interview-Based Impression of Change-Plus (rho=0.14).  Caregiv			_	"			-
Questionnaire in relation to the patient's engagement)    CC=0.58.							-
CPBQ): engagement   Also, the Rasch analysis of the 10-item CAT scale showed good overall fit suggesting a single   Scale 2: Two "factors" from the EFA analysis (not labeled) (Rasch analysis suggested a unidimensional (one-factor) (CAT)   Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).				, ,	thresholds not provided)		
Scale 2: Two "factors" from the EFA analysis (not labeled) (Rasch analysis suggested a unidimensional) construct.  CATE (CAT)  LOPEZ & ADRD Guarino  Scale 2: Two "factors" from the EFA analysis (not labeled) (Rasch analysis suggested a unidimensional) construct.  Concurrent validity was demonstrated by significant Spearman's rank correlations (p<0.001) between the CAT and the NPI (rho=0.35), the Severe Impairment Battery (rho=-0.19), the Alzheimer's Disease unidimensional (one-factor) (CAT)  Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Surrogate Guarino  Self-efficacy for decision making making  Self-efficacy for decision instrument's credibility, accuracy, and relevance as a measure of self-efficacy for surrogate decision  Rasch model) estimate for the full scale=0.83.)  Rasch model) estimate for the full scale=0.83.)				•			
Caregivers' analysis (not labeled) (Rasch Assessment of Themselves (CAT) construct.)  Lopez & ADRD Surrogate Guarino Making  ADRD Surrogate Guarino  Caregivers' analysis (not labeled) (Rasch analysis suggested a unidimensional (one-factor) (CAT) construct.)  Concurrent validity was demonstrated by significant Spearman's rank correlations (ρ<0.001) between the CAT and the NPI (rho=0.35), the Severe Impairment Battery (rho=-0.19), the Alzheimer's Disease unidimensional (one-factor) (Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Self-efficacy for decision making making A-point Likert scale instrument's credibility, accuracy, and relevance as a measure of self-efficacy for surrogate decision scale=0.83.)			'	,		,	
Assessment of Themselves (CAT) construct.)  Lopez & ADRD Surrogate Guarino  Assessment of Themselves (CAT)  Decision Making  Assessment of Themselves unidimensional (one-factor) construct. Self-efficacy for decision making  The CAT and the NPI (rho=0.35), the Severe Impairment Battery (rho=-0.19), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=-0.24), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Experiment Battery (rho=-0.19), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=-0.24), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Experiment Battery (rho=-0.19), the Alzheimer's Disease Cooperative Study-ADL Scale (rho=-0.24), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Experiment Study-ADL Scale (rho=-0.24), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Experiment Study-ADL Scale (rho=-0.24), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).  Experiment Study-ADL Scale (rho=-0.24), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.25), and the Functional Assessment Staging Tool (rho=0.14).  Experiment Study-ADL Scale (rho=-0.24), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.25), and the Functional Assessment Staging Tool (rho=0.14).  Experiment Study-ADL Scale (rho=-0.24), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.25), and the Functional Assessment Staging Tool (rho=0.14).  Experiment Study-ADL Scale (rho=-0.24), the Clinician's Interview-Based Impression of Change-Plus Caregiver Input (rho=0.25), and the Functional Assess						,	
Themselves unidimensional (one-factor) (CAT) (C						, , ,	scale=0.83.)
CAT)   Caregiver Input (rho=0.23), and the Functional Assessment Staging Tool (rho=0.14).							
Lopez & ADRDSurrogateSelf-efficacy for decision5 items,Face/content validitywas established by three expert Gerontological nurses who reported on theCronbach's α, full scale = 0.87GuarinoDecision Making making4-point Likert scaleinstrument's credibility, accuracy, and relevance as a measure of self-efficacy for surrogate decision							
Guarino Decision Making making 4-point Likert scale instrument's credibility, accuracy, and relevance as a measure of self-efficacy for surrogate decision	1 0	4 D C C	, ,	,	C :+		Creationable of full participation 2.27
		AUKU	_	•	· ·	, , ,	<u>cronpach's α, tuil scale</u> =0.87
[2013]**   peii-επicacy   Une factor: Seif-efficacy   (ranging from 1=Strongly   making. The reliability of agreement between the three experts was assessed with Fields' kappa				<u> </u>			
	(2013)	<u> </u>	реп-Епісасу	One ractor: Self-efficacy	ranging from 1=Strongly	linaking. The reliability of agreement between the three experts was assessed with Fielss' kappa	

		Scale (SDM-SES)		disagree to 4=Strongly	coefficient (Fleiss' kappa=0.90).	
United States		Scale (SDIVI-SES)		agree)	The structural validity of the scale was established through CFA of a hypothesized single underlying	
Officed States				agree	latent factor model for self-efficacy for decision making explaining the set of observed items. As	
					expected, CFA produced a single-factor (unidimensional) model with factor loadings ranging from 0.63	
					to 0.86. The model goodness-of-fit measures were acceptable (CFI=0.99; TLI=0.98).	
Bekhet &	ADRD	Depressive	Depressive cognitions	8 items,		Cronbach's α, full scale =0.88.
Zauszniewski	AUNU	•	One factor: Depressive	6-point Likert scale	The current study examined the structural validity of the scale with a PCA in a sample of ADRD CGs that	Cionbacii s u, iun scale -0.88.
(2013) <sup>86</sup>		(DPS)	cognitions	(ranging 0=Strongly	resulted in two factors/components. Authors follow-up with a CFA that produced a single factor	
(2013)		(DF3)	Cognitions	disagree to 5=Strongly	explaining 55.99% of the variance. This solution confirmed previous findings using the scale.	
United States				agree)	The <u>concurrent validity</u> was assessed through an expected positive Pearson correlation between DPS	
Officed States				agree	scores and Caregiver burden ( $r$ =0.40, $p$ <.001) measured by the ZBI and a significant negative correlation	
					with resourcefulness ( $r = -0.54$ , $p < .001$ ) as measured by the Resourcefulness Scale.	
M/illes of	ADRD	Cairitual Cuanart	Darasiyad spiritual support	12 itams		Cranbachis er full scala -0.074
			Perceived spiritual support One Factor:	12 items, 4-point Likert scale	The <u>structural validity</u> was demonstrated by an EFA with Varimax rotation that yielded a single-factor structure explaining 79% of the variance by a rotated Varimax solution.	<u>Cronbach's α, full scale</u> =0.974
al.(2013) <sup>89</sup> ,		Scale (SSS)		•	· · ·	Calit half raliability was actimated by
			(Items measure the use of	(1=Strongly disagree to		Split-half reliability was estimated by
United States			1 ' ' '	4=Strongly agree)		Guttman's coefficient showing a
			of coping.)		Resilience Scale (r=0.25, p<0.01). Validity was also supported by a significant negative correlation	strong correlation between two
					between SSS scores and the Emotion-Focused subscale of the CITS measure (r=-0.12, p<0.01). However,	
					SSS scores were not significantly correlated with the CITS's Avoidance-Focused subscale.	(Guttman's split-half reliability=0.940).
	ADRD	Impact of	Caregiver burden	12-items,	· · · · · · · · · · · · · · · · · · ·	<u>Cronbach's α, full scale</u> =0.927.
(2014) <sup>92</sup>		Alzheimer's	One factor	5-point Likert scale	generation being informed by reviewing the literature and identifying previous measures on AD	
		Disease on	(Items cover the following			Test-retest reliability (4-week interval)
United States		Caregiver	domains: Caregiver burden	to 4=Extremely)	experience of caring for a patient with AD and to conduct a cognitive debriefing of an initial 9-item draft	
			across emotional, physical,			subgroup of AD CGs (N=50). The ICC
		(IADCQ)	social, financial, sleep, and			was moderate (0.68).
			time impact)		The structural validity of the 12-item IADCQ was assessed through a CFA that resulted in a final one-	
					factor (unidimensional) solution that provided acceptable goodness-of-fit indexes (e.g., GFI=0.934;	
					RMSEA=0.076; CFI=0.944; and SRMR = 0.040).	
					Concurrent validity was assessed through "moderate to large" Pearson's correlations between IADCQ	
					scores and the Short Form-12 Health Survey (SF-12: V2) composite scores scales: Physical health (r= -	
					0.26, p < 0.001) and Mental health (r= -0.58, P < 0.001). Pearson's correlations between IADCQ scores	
					and other subscales from the SF-12: V2 were also "moderate to large" ranging from -0.20 to -0.57.	
Solberg et al.	ADRD	_	•	13 items,	Content validity. Authors developed a 32-item pool based on a literature review of the stress	Cronbach's α estimate for the 13-item
(2014) <sup>96</sup>		Impact Scale	, ,	3-point Likert type scale	experienced by caregivers for older adults in general. Items were adapted to reflect the impact of the	<u>scale</u> =0.74.
		(CGQ-13)	One factor: Impact of stress	with varying labels.	stress on adult children who were primary caregivers for their demented parents. (Adult children	
United States					caregivers were the primary focus of this study.)	
					The <u>structural validity</u> of the CGQ-13 scale was established via EFA with Oblique rotation and a scree	
					plot to determine the optimal number of factors. After item deletions due to low factor loading, the	
					scale was reduced to 13 items with high loadings on a single factor explaining 50% of the total variance.	
		Partnering for	CG activation	32-items,		Cronbach's $\alpha$ , full scale = 0.95
(2015) <sup>100</sup>			(CGs' knowledge and skills in		an initial item pool of 86 questions. Experts were asked to reflect on the items they considered	
		Living with	health care management of	(ranging from 1=Disagree	important for engaging CGs in patients' health care management and to identify skills that CGs must	Pearson's correlation coefficient was
United States		Chronic Illness:	persons with dementia and	, , ,	develop to support optimal health care. Cognitive interviewing was also conducted with 35 primary CGs.	
		Dementia (PBH-	the ability to meet their own	completely; with an	As a result of this step, a 35-item scale (23 "knowledge" and 12 "skills" items) emerged.	reliability (two-week interval) of the
		LCI: D)	needs.)	additional response	<u> </u>	scale scores in a sample of 79
			Six factors:	option: 0=Not my	rotation explaining 93.8% of the total variance. A scree plot confirmed a 7-component/factor underlying	participants (r = 0.76).
			(1) Understanding dementia	responsibility)	structure. Instead of fitting a multidimensional model, authors conduct a unidimensional Rasch analysis	
			(2) Recognizing and		with the initial 35-item pool. Despite the small sample size (N=130) and the underlying multi-	
			anticipating symptoms and		dimensional structure found in the previous step, most items showed acceptable fit statistics under the	
			challenges (3) Managing		unidimensional Rasch model. After eliminating 3 items due to poor performance in the Rasch analysis,	
			care patient's medications		the authors present the final scale as a "six-factor" 32-item scale.	
			(4) Managing day-to-day		Concurrent validity was established through significant Pearson's correlations (p-values < 0.05) between	
			symptoms and challenges		total scores on the PBH-LCI: D and scores on (a) Preparedness for Caregiving(r=0.69), (b) Global	
			(5) Recognizing sudden		Caregiving Self-Competence(r=0.41), (c) Global Caregiving Self-Confidence (r=0.43), and (c) the "mental	
			changes in patients' health		health component summary" obtained from the SF-12 (r=0.35). Scores on the PBH-LCI: D were	
1			(6) Utilizing health services		negatively correlated with anxiety measured by the General Anxiety Disorder Assessment (r=-0.33).	
Ī	1		and managing sudden		(Sample sizes used in the reported correlations ranged from N=52 to N=130).	

	1	Γ		1		
			changes in person's self-care			
Powers & Whitlach (2016) <sup>103</sup> United States		Justifications for Caregiving Scale (CJCS)	Cultural expectations and reasons for providing care (as a function of beliefs and norms about the caregiving role) Two factors: (1) Reciprocity (making a family contribution as a motivation for caregiving); (2) Duty (caregiving as a sense of duty or obligation)	3=Somewhat agree, 4=Strongly agree)	study reports the detailed psychometric properties of the scale in a diverse sample of <a href="mailto:dementia">dementia CGs</a> . <a href="mailto:structural validity">Structural validity</a> was assessed by PCA to extract the components/factors and Varimax rotation to facilitate the interpretation of item loadings. This analysis was conducted for the full sample and separately for the White and African American subsamples. The PCA analysis in the full sample produced a two-component/factor solution (labeled "Reciprocity" and "Duty") explaining 60% of the total variance. The pattern of loadings, however, <a href="mailto:differed">differed</a> across the White and African American subsamples suggesting lack of measurement invariance and the need to conduct formal invariance tests to meaningfully compare results between groups.	Note: No estimates per subscale (2-factors) were provided for the total
					research is needed regarding the measurement invariance of the scale across subgroups.	
Piersol et al. (2016) <sup>106</sup> United States	ADRD	Capacity Card Sort (FCCS)	CG appraisal of patient functional capacity (CG estimation/appraisal of patient's "function" regardless of the level of cognitive impairment)	from lowest Level 1 (automatic actions) to highest Level 6 (planned actions). The six cards describe an individual's ability to perform the daily activity of "washing self." Each card maps to a range of high/low modes	Content validity was assessed by seven experts (occupational therapists) who reviewed the original set of 12 cards and identified the intended Allen cognitive level of each card. Based on the level of accuracy achieved by raters the cards were collapsed into a final set of six cards and another group of five experts reviewed the cards achieving 100% accuracy. Three independent groups of CGs (N=72) also reviewed the final set of six cards for level of accuracy in terms of cognitive level and mode, level of difficulty, and clarity.  Concurrent validity was examined estimating the Spearman's rank correlation between the score on the activities of daily living (ADL) index of the Caregiver Appraisal of Function and Upset (CAFU) scale and	(90.3%) with the lowest level of function, next highest (86.1%) with the highest level of function, and less with the middle levels (74% and 76.4%).
Sadak et al.	ADRD	Kingston	CG stress	10 items,	Content/face validity was addressed briefly by the authors in the website description of the scale <sup>111</sup>	Cronbach's α, full scale =0.88.
(2017) <sup>110</sup> United States		Caregiver Stress Scale (KCSS)	Three factors: (1) Personal-/Caregiving- related stress; (2) Family- related stress; (3) Financial stress	5-point Likert scale (ranging from 1=no stress to 5=extreme stress)	indicating that content validity the KCSS was established by examining the scale questions and determining that they addressed the characteristics of caregiver stress.  Structural validity was established using a PCA that yielded a three-component/factor solution explaining 71% of the total variance. The three components/factors mapped on to a priory identified "domains" labeled as: Personal/caregiving-related stress, Family-related stress, and Financial issues.  Concurrent validity. Scores from subsamples completing the General Anxiety Disorder (N=51) scale and Patient Health Questionnaire (N=52) were significantly (p-values < 0.001) and moderately correlated with KCSS scores (r=0.69, 0.57, respectively).	Cronbach's α by subscales: Caregiving (α=0.885); Family (α=0.871) Financial (1 item, n/a) Test-retest reliability (two-week interval) in a subsample (N=78): Pearson's r=0.88.
Piggott et al. (2017) <sup>112</sup> United States		Confidence in Sign/Symptom Management (CCSM) Scale	CG self-efficacy (confidence) in sign/symptom management; CG role strain Four factors: (1) Knowledge of signs/symptoms; (2) Management of cognitive signs/symptoms; (3) Management of medical signs/symptoms; (4) General medication management	5-point Likert scale (ranging from 1=Not at all true/confident to 5=Extremely true/confident)	initial 37-item bank. They were also asked to provide recommendations of additional questions concerning their relative's medical problems or about their own self-efficacy not measured in the current scale. Further revisions reduced the original scale to 26 items.  The <u>structural validity</u> of the 26-item scale was established through EFA with a Promax (oblique) factor rotation followed by the examination of the scree plot to determine scale dimensionality. After eliminating an item, the final 25-item CCSM scale produced a four-factor solution.  Concurrent validity of the CCSM scale was assessed by Pearson's correlations with 3 widely used CG measures: (1) the ZBI-role strain (r=-0.36, p <0.001) and the ZBI-personal strain (r=-0.14, p=0.06); (2) the Generalized Anxiety Disorder-7-item scale (r=-0.12, p=0.09); and (3) the 9-item Patient Health Questionnaire-Depression (r=-0.06, p=0.43). CGs with less role strain reported more confidence in all subscales (correlations ranged from 0.37 (p < 0.001) for general medical management to 0.15 (p=0.042) for knowledge about signs/symptoms. The association between the caregiver's self-report of medical training and CCSM scores was also significant (r=0.26, p < 0.001).	Test-retest reliability (2-day interval) was assessed with N=17 CGs using

Woisma	ın de AI		Ctigma Impact	Stirms	24 itans		ICC=0.56); Management of cognitive signs/ symptoms (r=0.87, ICC=0.82); Management of medical signs/ symptoms (r=0.78, ICC=0.78); General medication management (r=0.95, ICC=0.94) Cronbach's α, full scale =0.93.
Mamani				Stigma Four domains:	24-items, 4-point Likert scale	The <u>structural validity</u> of the scale is not established as part of the current study with dementia CGs. Authors relied on the 4-domains of SIS defined by Burgener & Berger (2008) <sup>122</sup> using an adapted version	Cronbach's $\alpha$ , full scale =0.93.
al. (2018			' '	(1) Social Rejection; (2)		of the original scale in a <u>different population</u> of CGs. <u>Content validity</u> was also examined in the adapted	
ui. (2010	,			, ,		version. Although the objective of the current study was not to establish the validity of the SIS scale in a	
United 9	States			• • • • • • • • • • • • • • • • • • • •		sample of dementia CGs, the study provides evidence of the concurrent validity and reliability of SIS	
				Social Isolation	<i>o</i> ,	among dementia CGs. Authors hypothesize an association between SIS measures and constructs	
						measured by Expressed emotion assessed using the 20-item Family Questionnaire (FQ). FQ also has two	
						subscales: Emotional Over involvement (EOI) and Criticism. As hypothesized, greater CG stigma was	
						positively associated with Criticism (r=0.372, p < 0.001) and EOI (r= 0.398, p < 0.001). EE total scores	
						(i.e., the sum of the Criticism and EOI subscales) were also significantly correlated with stigma (SIS)	
D	-1 -1 05	200	D	CCi CC ll	2411	scores (r= 0.434, p < 0.01).	Mar Danaldla Garda Gillarda O O O
(2018)12	et al. Al		Dementia Burden Scale –		34 items,	The DBS-CG scales was developed by combining 34 items from existing scales.  The structural validity for the 34-item scale was established through two alternative CFA models: a 3-	McDonald's ω for the full scale=0.93.
(2018)						factor model and a bifactor model (one general factor and 3-specific factors) using items from three	
United 9	States			Distress caused to the CG by		existing scales: The Modified Caregiver Strain Index (MCSI), the NPI Questionnaire-Distress scale, and	
				the patient's behavioral		the Patient Health Questionnaire (PHQ-9). The resulting models fit the data well but the bifactor model	
				symptoms; (3) Depressive	"Not distressing at all" to	produced a slightly better fit: (RMSEA=0.05, CFI 0.95). The score in the general factor represented	
				symptoms	"Extreme or very severe";	9	
						Responsiveness-Minimal important differences estimates of the amount of clinically relevant change on	
					every day"	the scale ranged from 4 to 5 points (effect sizes associated with each of these differences were "small":	
Davis st	- L A F	000	C:It After	Cuilt and ambinalance	10 :	0.20–0.49).	Create ship of full scale 0.02
Davis et (2019) <sup>12</sup>				Guilt and ambivalence following nursing home		Although no formal statements on <u>content validity</u> are made, authors developed scale items through information obtained from focus groups and a literature review of the emotional aspects of placement.	<u>Cronbach's α, full scale</u> =0.92.
(2013)				placement	•	The focus groups consisted of a study clinician attending caregiver support groups run by the	
United S	States		-	<u> </u>		Alzheimer's disease association to explore CGs' feelings regarding nursing home placement. An initial	
				reflecting guilt associated	, ,	sample of 46 items was generated using this method.	
				with making the decision to		Using an initial 46-item pool, the scale's <u>structural validity</u> was assessed via EFA with PAF to extract	
				place		factors and Varimax rotation to explore factor loadings. After several EFA iterations and refinements, a	
						10-item GAP-Q scale produced a single underlying factor (1-factor solution) as the best fitting model.	
						Concurrent validity was evaluated in a subset of the sample (N=53) using Pearson correlations between	
						the GAP-Q scores and concurrent measures of (a) depression using the CES-D ( $r=0.53$ , $p<0.001$ ), (b) CG burden using the ZBI ( $r=0.48$ , $p<0.001$ ), (c) conflict with staff using the Interpersonal conflict scale (ICS)	
						(r=0.47, p <0.001), and (d) "wellbeing" using the short form health survey (SF-36) (r= -0.30, p <0.05).	
Wynn &	. М	1ixed	The	Frontotemporal dementia	18 items,	No formal statements on <u>content validity</u> are presented. However, authors reported reviewing the	CG Sample:
Carpent	er	ļ		•		literature to ensure item content relevance and coverage. The research team also reviewed an initial	Cronbach's α, full scale =0.846.
(2020)13	37			One factor (Knowledge of		24-item pool, removed items with overlapping content, and rewrote items for clarity achieving a	Split-half reliability (Spearman-Brown)
						twelfth-grade reading level.	= 0.814.
United S	States		'	(1) Risk factors		No factor analysis to examine the underlying factor structure of the scale is reported, but authors state	
						that the scale "measured a unidimensional construct of knowledge about FTD". In the CG sample, convergent validity was demonstrated by correlating FTDKS and level of care provided	Cronbach's α, full scale = 0.704.
				(3) Disease course (4) Caregiving		to people with FTD (Pearson's r=0.231, p < 0.05). In contrast to expectation, scores on the FTDKS were	
				( ) Caregiving		not correlated with the number of people with FTD known (r=0.179, ns).	0.723.
Van Hou	utven AD	DRD	Caregiver	Perception of support. CGs	12 items,		McDonald's ω by subscales:
et al.				perceptions of support from	4-point Likert scale		Communication (ω=0.90)
(2020)13	88			the patient's health care	r	health care encounters (e.g., Consumer Assessment of Healthcare Providers and Systems [CAHPS]), and	Capacity (ω=0.94)
			Communication			an organizing framework of CG skills. No further pilot testing steps are provided.	
United S	states			communication experiences		The <u>structural validity</u> of the CAPACITY scale was established by CFA. A model with a two-factor	
1				with the team. Two factors:		structure (with factors labeled as "Capacity/preferences" and "Communication") was the best fitting	
		I.	(CADACITY)	(1) Canacity/proformance (2)		model Goodness of fit indices were acceptable (e.g. PMSEA-0.00E) CEI-0.073; and TII-0.067)	
				<ol> <li>Capacity/preferences;</li> <li>Communication</li> </ol>		model. Goodness-of-fit indices were acceptable (e.g., RMSEA=0.085; CFI=0.973; and TLI=0.967).	

Galvin et al. (2020) <sup>142</sup> United State		Negative Appraisals of Caregiving (PANAC) Scale	Two factors: (1) Positive Appraisals (PAs) (2) Negative Appraisals (NAs)	(ranging from 1=Strongly disagree to 5=Strongly agree	The <u>structural validity</u> of the scale was determined by EFA with principal components as the factor extraction method and Varimax rotation that produced a 2-factor/component solution explaining 46.7% of the cumulative variance. <u>Concurrent/discriminant validity</u> was assessed using Pearson's correlation coefficients between PANAC PA subscale scores and a) the Applied Mindfulness Process Scale, AMPS (r=0.31, p=0.001), b) CG burden, ZBI (0.014, ns), c) CG depression scores, PHQ4 (0.026, ns), and (d) a 4-item CG self-efficacy measure developed by the authors (0.073, ns). Negative Appraisal (NAs) of caregiving were associated with AMPS-low positive emotional regulation (r=-0.25, p=0.013), lower self-efficacy (r=-0.55, p < 0.001), higher ZBI scores (r=0.52, p < 0.001), and greater CG depression (r-0.37, p < 0.001).	
Teresi et al. (2020) <sup>153</sup> United State	ADRD S	Perceived Stress Scale (PSS)	One factor: Stress	10-item, 5-point Likert scale (0=Never, 1=Almost never, 2=Sometimes, 3=Fairly often, 4=Very often)	model with polychoric correlations. PCA was conducted <u>separately</u> for the total sample and selected subgroups: age, education, and language of the test (Spanish-English). The PCA suggested a one-factor/component model explaining 54% of the variance for the total sample and ranged from 50% to 57% for subgroups. The bifactor CFA confirmed the unidimensionality of the scale producing goodness-of fit indices within acceptable thresholds (RMSEA=0.044, CFI=0.996).  To provide validity evidence based on the internal structure of the PSS scale, <u>differential item functioning</u> (DIF) was examined for <i>age</i> , <i>education</i> , and <i>language</i> using the graded response IRT model. In general, the magnitude and impact of DIF were minimal across the groups examined.	Cronbach's α ordinal estimate, full scale=0.902, McDonald's $\omega$ =0.904, and the bifactor model explained common variance, ECV=68.34. IRT-based reliability measures were examined at selected points along the underlying latent continuum (attribute levels). The average reliability estimate for the total sample was 0.89 and ranged from 0.88 to 0.90 for subgroups. Test-retest reliability (6-month interval) examined over three followup waves (with samples N=343, 301, and 219). McDonald's $\omega$ estimates were about 0.90 across waves.
Thompson e al. (2020) <sup>154</sup> United State		Incompetence— Dementia Scale (FOI-D)	·	7-point Likert scale (1=Not at all concerned to	dementia caregivers for clarity and suitability. Based on the feedback, seven items were added and a preliminary 87-item scale was field-tested.  The <u>structural validity</u> of the scale was established by iterative EFAs, using ML as the factor extraction approach, followed by CFAs to cross-validate the identified factors structure. The iterative analyses resulted in a final 58-item scale that supported a 3-factor structure. Goodness-of-fit indices for the CFA	
Rose et al. (2021) <sup>157</sup> United State	ADRD S	Life in Dementia (FQOL-D) scale.	caregiving on family quality- of-life Four factors:	41 items, 5-point Likert-type scale (ranging from 1=Very dissatisfied to 5=Very satisfied)	Face validity was established by 2 persons with early stage ADRD and six family CGs who provided input regarding the clarity, readability, and content of the items included on the proposed FQQL-D instrument. Content validity was established by a panel of experts in ADRD research and care from across the United States who reviewed items for clarity of expression. A Delphi method was employed to identify important factors of family quality of life in dementia given 5 previously identified domains and preliminary items. Items were retained by panel consensus. Experts were given the opportunity to write in additional items not originally included. The final item pool comprised 43 items.  Structural validity was assessed by factor analysis with PCA as the extraction method and Varimax rotation to increase interpretability of the factors/components. The PCA provided support for a 4-factor solution that explained 52% of the variance in the scale items.  Concurrent validity was established by correlating the FQOL-D scale with three scales: 1) the "Family Resource", 2) the Family "Adaptation, Partnership, Growth, Affection, Resolve" (APGAR), and 3) the "Surrogate Decision Making Self-Efficacy scales". Increased FQOL-D scores were associated with higher scores in each of these scales. Pearson's correlations ranged from 0.39 to 0.46 (p-values < 0.01).	

		<u>Scale II</u> . Other	behavior employed by other relatives (e.g., siblings, husband) Two factors: (1) Accusations of harming the patient; (2) Shifting	scored on two scales:	dimensionality, and a CFA. The analyses supported a two-factor structure. Goodness-of-fit indices for the CFA model were acceptable (e.g., RMSEA=0.01; CFI=0.99; and TLI=0.99).  Concurrent validity was established by Pearson correlations between the Caregiver Guilt Questionnaire (CGQ) developed by Losada et al. (2010) <sup>74</sup> and the two subscales. Only the "shifting responsibility onto	The Cronbach's $\alpha$ , full scale =0.78. Cronbach's $\alpha$ by subscales: Accusations of harming the care recipient. ( $\alpha$ =0.81) Shifting responsibility onto the CG ( $\alpha$ =0.80)
Park et al.	ADRD	Pre-Loss Grief-	responsibility onto the CG Pre-loss grief	10 items,	The PG-12 was originally developed for non-AD carers and contained 12 items. <sup>169</sup> The current study	The Cronbach's α, full scale = 0.89
(2022) <sup>168</sup>		10-Dementia	One factor: Grief symptoms	5-point Likert scale	reduces, adapts, and validates the scale with a sample of dementia CGs.	
		(PG-10-D)		(from 1=Almost never to	The <u>structural validity</u> of the scale was assessed through iterative CFA producing a final one-factor	
United States				5=Always)	(unidimensional) model with 10 items and factor loadings ranging from 0.53 to 0.85. Goodness-of-fit	
					indices were within acceptable ranges (e.g., RMSEA=0.06; CFI=0.97; and TLI=0.96).	
Kuzmik et al.	ADRD	Modified	Caregiving strain	13 items,	Content validity was assessed by the original developer of CSI. <sup>181</sup> The scale was later modified by	This study did not report reliability
(2023)180		Caregiver Strain		3-point scale (0=No,	, , ,	measures for the sample dementia
		Index (MCSI)		1=Yes, sometimes, 3=Yes,	Structural validity. CFA was performed to test the one- and two-factor models of the MCSI identified in	I
United States				on a regular basis)	prior studies. The two-factor model provided a better fit. Factors were labeled: individual experiences of	
			on the CG's life		burden and repercussions on the CG's life. Reported "goodness-of-fit" measures were within acceptable	
			Factors include the			and Travis (2003) <sup>182</sup> using the MCSI
			following "domains:"		, , , , , , , , , , , , , , , , , , , ,	reported a Cronbach's α of 0.90 and a
			financial, physical,			test–retest (2-week interval) reliability
			psychological, social, and			coefficient of 0.88. However, these
			personal.			estimates were obtained from a <i>mixed</i>
					Measurement Invariance. Tests of measurement invariance by race (configural, metric, and scalar) were	I
					•	unspecified disease status.
					multigroup CFA model produced results confirming measurement invariance by race.	

Note: AD = Alzheimer's disease; ADRD = Alzheimer's disease and related dementias; ADL = Activities of Daily Living; AGFI = adjusted goodness-of-fit index; AVE = average variance extracted. A recommended threshold for convergent validity is an AVE > 0.50; CG = Caregiver; CATPCA = categorical principal component analysis; CES-D = Center for Epidemiological Studies Depression Scale; CFA = confirmatory factor analysis; CFI = comparative fit index; CR = composite reliability. A recommended threshold for convergent validity is a CR > 0.70; CVI = content validity index; <sup>191</sup> EFA = exploratory factor analysis; GFI = goodness of fit index; Hamilton Depression Rating Scale = HAM-D; Hospital and Anxiety Depression Scale = HADS; IADL = instrumental activities of daily living; ICC = Intra-class correlation coefficient; IFI = incremental fit index; IRT = item response theory; LSNS= Lubben Social Network Scale; ML = maximum likelihood; MLE = maximum likelihood estimation; MMSE = Mini-Mental State Examination; NPI = Neuropsychiatric Inventory; NFI = Normed Fit Index; NNFI = non-normed fit index; PAF = principal axis factoring; PCA = principal components analysis; POMS= Profile of Mood States; RMPBC = Revised Memory and Behavior Problems Checklist; RMSEA = root mean square error of approximation; SF-36 = Short form 36 Health Survey; SRMR = standardized root-mean-square residual; TLI = Tucker-Lewis Index; ZBI = Zarit Burden Interview; PSI = person separation index.<sup>192</sup> PSI values above 0.70 indicate good to excellent reliability in differentiating persons along the measured trait. Proposed rule of thumb thresholds for ICCs are: between 0.50 and 0.75 (moderate); ≥ 0.75 (good), and ≥ 0.90 (excellent).<sup>193</sup> Generally accepted threshold for "good" Cronbach's α test of reliability is considered to be ≥ 0.70. Responsiveness (longitudinal validity) refers to the ability of an instrument to detect clinically important changes over time.<sup>194</sup> Measures such as minimal important change (MIC), smallest detectable change (SDC), effect si